

Skin in the Game: Personal Stock Holdings and Investors' Response to Stock Analysis on Social Media

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ABSTRACT:

Motivated by concerns that financial positions present a conflict of interest that impairs an analyst's objectivity, we examine investor perceptions of the financial positions of non-professional analysts (hereafter NPAs) providing stock analysis on the social media outlet SeekingAlpha (hereafter SA) and offer two primary findings. First, NPA positions contribute directly to short-window returns surrounding the article's publication, holding constant the information in the article (i.e., tone, length, rigor, numerical content, etc.) as well as contemporaneously issued news (i.e., from professional analysts, managers, and the business press). Economically, disclosure of a long position by an NPA corresponds to a positive 2-day return of 0.4 percent, while disclosure of a short position corresponds to a -1.0 percent return over the same period. These findings suggest that an NPAs stock positions convey information to investors. Second, contrary to concerns that stock positions are associated with biased analysis, we find no evidence that NPA positions reduce investor responses to the tone of the article. In fact, our evidence suggests that holding a position *magnifies* investor responses to both positive and negative tone, and these effects seem to be primarily driven by tone that is contrary to the NPA's stock position. Overall, our findings suggest that, contrary to regulators' concerns, stock positions of non-professional analysts do not appear to decrease the credibility and informativeness of their analyses.

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1. Introduction

A large body of research establishes that professional financial analysts play a valuable role in the capital markets by providing both new information and interpreting previously released information (Womack 1996; Asquith, Mikhail, and Au 2005; Bradshaw 2011; Bradshaw, Wang, and Zhou 2017; Brown, Call, Clement, and Sharp 2015, 2016; among others). In addition, over the last decade, access to financial information has exponentially increased, leading to the proliferation of nonprofessional analysts on social media (Chen, De, Hu, and Hwang 2014; Drake, Thornock, and Twedt 2017) that has led some in the industry to question whether the role of professional financial analysts will eventually become obsolete (Dediu 2011; Chernova 2014). However, a challenge to the use of nonprofessional analysis by investors is the lack of regulation: unlike professional analysts, non-professional analysts (hereafter, NPAs) face little regulatory oversight, leading to the potential for market manipulation.¹

We examine the information conveyed by analysis posted to the social media site SeekingAlpha (seekingalpha.com, hereafter, SA)² to evaluate whether perceived credibility (i.e., investor response to analysis) varies with an NPA's personal stock position in firms about which they write. On the one hand, practitioners and regulators have expressed concern that stock positions may impair the objectivity of financial statement analysis. For example, the Securities and Exchange Commission (SEC) argues that stock positions “can create pressure on...independence and objectivity,” although “the existence of these relationships does not necessarily mean...bias” (SEC 2016). On the other hand, NPAs with personal stock positions have

¹ Unlike professional analysts or firm insiders, NPAs publishing on sites like SeekingAlpha do not face legally enforced “blackout periods” that limit trading activity around the publication of their reports.

² There are a number of social media venues considered by prior research such as Motley Fool, Estimize, StockTwits, Yahoo! Finance, and online stock message boards. We focus on SA because (1) it requires contributors to provide written, edited analysis (suggesting a certain level of rigor and sentiment, or tone), and (2) it requires contributors to disclose whether they have a financial position in the firms about which they write. Chen, De, Hu, and Hwang (2014) find that SA content represents value relevant information. Additionally, Drake, Thornock, and Twedt (2017) identify SA as a credible internet-based information intermediary.

a vested interest in the stocks about which they write, which may enhance the quality, rigor, and timeliness of their analysis.

We examine two specific research questions. First, does an NPA's financial position convey information incremental to the content of their analysis? If so, the direction of that position could reveal the NPA's private information to investors. On the other hand, if a financial position indicates bias on the part of the author then investors would not find the position to be incrementally informative. Second, do investors find the tone in an NPA's analysis to be more credible when the author has a financial position (i.e., the author has "skin in the game")? If holding financial positions increases the quality of the NPA's analysis, investors may respond more strongly. Conversely, if investors perceive NPAs with personal stakes to be biased, we may observe a weaker response to the analysis.

SeekingAlpha (SA) requires NPAs to include position disclosures in the articles they write. Using position disclosures from 104,952 SA articles from 2006 to 2015, we find that NPA positions contribute directly to short-window returns surrounding the article's publication, holding constant the information in the article (i.e., tone, length, rigor, numerical content, etc.) as well as contemporaneously issued news (i.e., from professional analysts, managers, and the business press). In terms of economic significance, we find that the disclosure of a long (short) position by an NPA corresponds to a 2-day return of 0.4 percent (-1.0 percent).³ These findings suggest that investors view a position disclosure as an information signal in its own right, presumably about the NPA's private information that is not included in the article.

³ Based on data provided by George Morarty, SeekingAlpha's editor, SA is used by both institutional and retail investors. His data suggests that the professional investors that use SA control \$15 trillion in managed assets (from both institutions and high-net-worth-clients), which compares to \$1.3 trillion in savings and investments of retail investors that use their service. Thus, if SA position disclosures are informative, users of the information have the purchasing power to significantly move stock prices.

Next, we find no evidence that NPA positions reduce investor responses to the tone of the article, again challenging the view that the effect of analyst stock positions is to produce biased analysis. Instead, our evidence suggests that NPA positions, if anything, make the analysis *more* credible. Specifically, we find that NPA positions appear to *magnify* investor responses to tone. Additional analysis suggests this result is primarily driven by tone contrary to an NPA's position (i.e., short positions magnify the response to positive tone and long positions magnify the response to negative tone). In that sense, investors find the NPA to be most credible when they provide information about a firm that goes against their personal financial interests. However, it is important to note that tone that is directionally consistent with an NPA's position is not discounted relative to analysis by NPAs with no positions, again inconsistent with financial positions inducing credibility-reducing bias.

An important caveat is that we are not aware of any mechanism through which SA obtains information about an NPA's investment portfolio, so they cannot audit an NPA's position disclosure. Therefore, it is possible that an author could intentionally mislead investors with their position disclosures. For that reason, we examine whether the contemporaneous reaction to a position disclosure is met with a subsequent reversal of that reaction. We find no evidence of a reversal of investors' initial responses, which is inconsistent with position disclosures being untruthful. In fact, we find that short positions are associated with continuing *negative* returns over the 60 days following the disclosure (i.e., a "drift" rather than a "reversal"). Taken together, these findings suggest that investors find position disclosures credible and useful for assessing the NPA's private information, and that NPA positions increase the credibility and informativeness of their analyses rather than constituting conflicts of interest.

We expand upon our primary findings through several additional analyses and robustness tests. Specifically, we perform several tests to mitigate the likelihood that our results are attributable to other major events occurring concurrent to each article’s publication. First, we restrict our sample to articles published in the early trading hours of the equity market. As we explain in Section 5.1, SA’s editorial process makes it virtually impossible for early-morning articles to be written about events occurring on the same day as the article’s publication. Results in this subsample are identical to those previously discussed, mitigating the concern that our results are due to investor reaction to an event other than the article release. Second, we test whether the reaction to an NPA’s position strengthens with article length, rigor, and numerical content. Consistent with this expectation, we show that the positive (negative) association between long (short) position disclosures and returns strengthens with the length of the article. We also find some evidence that numerical content strengthens investor reactions to disclosure of short positions. These results suggest an interactive effect between the effort put forth by the NPA and the information conveyed by their stock position. Third, we examine whether the first-time disclosure of a position is more informative than subsequent disclosures and find that the reaction to both short and long disclosures is significantly stronger the first time an NPA discloses a position.⁴

Our study provides several contributions to the accounting and finance literatures. First, we contribute to the literature on the informativeness of “crowd sourced,” peer-based advice (e.g., Chevalier and Mayzlin 2006; Liu 2006; Chen and Xie 2008; Zhu and Zhang 2010; Jame, Johnston, Markov, and Wolfe 2016; Drake, Thornock and Twedt 2017; Tang 2017). Through an examination

⁴ We expect first-time disclosures to be most informative because they indicate new information about an author’s position that may be value relevant. However, repeat disclosures indicate a continued commitment to the position and likely still provide information.

of SA articles, Chen, De, Hu, and Hwang (2014) take a first step towards addressing the question of whether “crowd sourced” financial statement analysis on social media sites conveys credible and value relevant information, or if instead such analysis represents “noise” or even an attempt to mislead. They find that, on average, these articles provide value relevant information that is incremental to traditional information sources such as the business press and professional financial analysts. Because SA is not directly regulated and NPAs, unlike financial journalists, lack established rules of conduct, these findings suggest that users of social media must find alternative mechanisms for assessing the credibility of information. Consistent with this supposition, Chen et al. also find that investors perceive the information in SA articles to be more credible when NPAs have an established track record of providing value relevant analysis. We identify a credibility-enhancing signal that can vary by NPA: investors perceive NPAs as more credible if they hold a position in the firm’s stock, thus aligning their personal incentives with either long or short traders. Furthermore, we find no evidence that NPAs exploit investors’ trust, on average, despite no direct enforcement mechanism to ensure that they report their positions honestly, suggesting that social and reputational pressures are largely sufficient to induce honest reporting.

Second, we contribute to prior research on analysts’ conflicts of interest. Much of this research finds that analysts have significant conflicts of interest placed on them by their firms (e.g., Lin and McNichols 1998; Michaely and Womack 1999; Dechow, Hutton, and Sloan 2000; Daniel, Hirshleifer, and Teoh 2002; Bradshaw, Richardson, and Sloan 2006; Ke and Yu 2006). In a review of this literature, Bradshaw (2011) notes that one of the most prevalent beliefs in the capital markets is that analysts’ behavior is dominated by conflicts of interest. Current SEC rules not only require analysts to disclose their positions, but also impose strict rules on the timing and nature of

those positions to mitigate conflicts of interest.⁵ Consistent with these regulations, Chan, Lin, Yu, and Zhao (2018) find that analysts with personal holdings provide recommendations perceived as more credible. However, the NPAs in our study are not bound by regulations on trading activity, freeing them to profit should they choose to mislead investors, yet our evidence also suggests that investors find NPAs more, not less, credible when they hold positions in the firms they follow. This finding could suggest either that holding a position in a stock does not by itself create a significant conflict of interest, or that disclosure of positions is sufficient to alleviate conflicts of interest. In either case, our study is of interest to regulators in evaluating what regulatory restrictions should be placed on financial analysts and other information intermediaries.

Third, we contribute to the literature on the role of the business press in financial markets by investigating how NPA financial positions affect investors' use of information. Prior research largely examines how the financial press contributes to a firm's information environment through the dissemination of information (Barber and Loeffler 1993; Huberman and Regev 2001; Busse and Green 2002; Tetlock 2007, 2010; Bushee, Core, Guay, and Hamm 2010; Engelberg and Parsons 2011; Dougal, Engelberg, Garcia, and Parsons 2012; Gurun and Butler 2012; Bradshaw, Wang, and Zhou 2015; Li 2015; Blankespoor, deHaan, and Zhu 2018). Collectively, this research suggests that the financial press plays an important role in the origination and dissemination of information but does not consider whether NPA positions in the firms they cover affect this process. We contribute to this research stream as the first to examine whether these stock positions impart incremental information to investors and whether they enhance or impair credibility.

⁵ The current analyst disclosure rules were initially developed by the New York Stock Exchange (NYSE) and the Financial Industry Regulatory Authority (FINRA), and then adopted by the SEC in 2002 (SEC 2016).

2. Background, Prior Literature, and Hypothesis Development

2.1 *Personal stock positions and credibility*

Whether personal holdings affect the credibility of opinions disseminated by non-professional analysts (NPAs) is an empirical question. Research identifies two primary sources of conflicts of interest for professional analysts that may impair the credibility of their work. Specifically, there are *firm-related conflicts* such as the generation of investment banking fees, trading commissions, trading gains/losses, etc. (e.g., Lin and McNichols 1998, Michaely and Womack 1999; Dechow, Hutton, and Sloan 2000; Daniel et al. 2002; Bradshaw, Richardson and Sloan 2006), as well as *personal conflicts* such as their compensation structure, long-term reputation, job security, need to ingratiate themselves to managers and powerful investor groups, and personal trading gains/losses (Ke and Yu 2006; Bradshaw 2011).

Regulators worry that stock positions of analysts could create an additional bias, impairing the credibility of their reports and recommendations (SEC 2016), and experimental research provides evidence consistent with this concern. Specifically, Taha and Petrocelli (2014) and Marley and Mellon (2015) directly investigate conflicts of interest arising from analysts' personal financial positions and find evidence that analysts with stock positions are *less* credible. However, both experiments use a single-period game in a laboratory setting, which removes any possible effects of analyst reputation. Relatedly, Bradshaw, Huang, and Tan (2014) find that, in an international setting, analysts' conflicts of interest contribute to forecast bias, and Chan et al. (2018) find that "analyst-owners", or analysts covering firms in which they own stock, provide upwardly biased price targets. However, Chan et al. (2018) also suggest that these same analysts issue recommendations that are perceived as more informative by investors.

There is also prior research on whether stock positions impact investors' perceptions of disclosures provided by other investors, such as firm-insiders (managers). To alleviate conflicts between managers and shareholders, 92 percent of firms adopt some type of policy regarding insider trading (i.e., "blackout periods") (Bettis, Coles, and Lemmon 2000), whereby insiders are prohibited from trading during the trading days surrounding an earnings announcement. Bettis et al. (2000) present evidence that "blackout periods" on manager trading reduce adverse selection costs, suggesting that when firms impose trading prohibitions on insiders, investors perceive the information as more credible. Furthermore, insiders are often privy to substantial value-relevant information not possessed by the market as a whole (Jaffe 1974; Seyhun 1986). As such, it seems natural that investors would want to observe their portfolio positions and changes in these positions. Indeed, recent studies suggest that timely information about manager purchases convey information to market participants (Fidrmuc, Goergen, and Renneboog 2006; Brochet 2010).

With respect to professional investors, the literature suggests that institutional investors try to withhold rather than provide private information. For instance, Agarwal, Jiang, Tang, and Yang (2013) find that when hedge funds ask the SEC to keep their ownership levels confidential, these positions are associated with information-sensitive events and higher information asymmetry. They conclude that stock positions of hedge funds convey information about their private information. Similarly, Aragon, Hertz, and Shi (2013) conclude that hedge fund managers seek confidentiality to protect proprietary information.

NPAs publishing on SA share attributes with many of these groups. They are very similar to professional analysts (because they publicly provide detailed analyses about firm value) and investors (because they describe themselves as active investors). While not insiders, SA contributors may also have an information advantage over other investors because of unique access

to management (Seeking Alpha 2017).⁶ However, there are also important differences between SA NPAs and these groups. SA NPAs *voluntarily* disclose their private information and face no trading restrictions (i.e., blackout periods) or enforcement (i.e., no portfolio audits). Thus, SA NPAs could either immediately trade out of a position after publishing an article or provide a false disclosure in an attempt to manipulate price (e.g., falsely disclose a short position prior to purchasing a stock). However, while strategies such as this are possible, they are unlikely to be sustainable in a multi-period setting without anonymity. We think it is more likely that NPAs provide credible private information in order to bolster their reputation in the investing community, providing opportunities to sell their analysis to others, and to accelerate price discovery for their stock positions (Pasquariello and Wang 2018; Ljungqvist and Qian 2016). In the latter case, personal stock holdings could enhance a NPA's credibility if it implies that the individual is confident enough in their information set to "put their money where their mouth is."

2.2 Social media and investor-sourced stock opinions from SeekingAlpha

Social media allows investors to supplement information from traditional sources by communicating directly with one another.⁷ Although the method of information sharing makes a difference, with some venues (e.g., internet stock message boards) seeming to produce mostly noise and confusion (Antweiler and Frank 2004; Das and Chen 2007), recent research suggests that social media can produce and disseminate value-relevant information. For instance, both Chen et al. (2014) and Jame, Johnston, Markov, and Wolfe (2016) find evidence that social media (i.e., SA and Estimize, respectively) communicates new information to the market. The latter study

⁶ According to SeekingAlpha's website, one benefit of being a SA contributor is "Access to Company Management" (<https://seekingalpha.com/page/become-a-seeking-alpha-contributor>). There, they mention that "[c]ompanies pay close attention to what is written about them on SA. Some companies also contribute via articles and comments. Many contributors have been given exclusive access to company executives to get their side of the story."

⁷ A related stream of literature examines how firm insiders, such as managers or employees, communicate with market participants via social media (e.g., Blankespoor, Miller, and White 2014; Hales, Moon, and Swenson 2018).

finds that “crowdsourced” earnings estimates are as accurate as professional analyst forecasts for some horizons, lending credence to speculation that the role of the paid, professional financial analyst might eventually become obsolete (Dediu 2011). Using Twitter, research similarly links aggregated sentiment to both future sales and earnings announcement news (Tang 2017; Bartov, Faurel, and Mohanram 2018). Finally, Drake, Thornock, and Twedt (2017) identify the internet as a new important “information intermediary” and suggest that content published on sites like SA improves price efficiency.

One of the largest social media platforms, SA, has become a popular venue for both professional and non-professional investors to share the results of their own analysis of financial securities. SA is rapidly becoming one of the most referenced sources for financial news and analysis. Investopedia.com ranks it third, behind only Google Finance and Yahoo! Finance, and users of “The top tens” rank SA first, ahead of both the Wall Street Journal and Financial Times.⁸ Citing Chen et al. (2014), the Wall Street Journal even speculates that NPAs publishing on sites like SA could replace professional financial analysts (Chernova 2014). SA reports an average of 7 million unique visitors per month and states its mission is to provide “opinion and analysis rather than news...written by investors...rather than journalists” (Seeking Alpha 2016). They publish an average of 200 to 250 articles per day, which, with their subscription base, corresponds to 200 million email or mobile alerts going out each month. SA does not generally solicit opinions or content, but does pay contributors based on the number of users accessing their content. Chen et al. (2014) suggest that the long form of SA articles, combined with the curation of content by SA’s editorial board, results in the identification of NPAs with something valuable to say and an opportunity for them to say it. Consistent with this suggestion, they find that the fraction of

⁸ See <http://www.investopedia.com/articles/investing/112514/top-sites-latest-stock-market-news.asp> and <http://www.thetoptens.com/financial-news-websites/>. Both sites accessed in Summer 2017.

negative words in a Seeking Alpha article is negatively associated with both stock returns over the following three months and subsequent earnings surprises.

In addition to its broad readership, SA is unique from other social media platforms in that the articles provide substantial, edited analysis, which may or may not include a formal “recommendation.”⁹ Platforms like Estimize (see Jame et al. 2016; Da and Huang 2017) provide an earnings estimate without any analysis. Platforms like stock message boards or Twitter allow any user to post information without quality control. SA articles, and in particular the long form articles we sample, provide in-depth analysis that is edited to ensure quality control.

In conclusion, prior research establishes that, on average, social media represents an important and emerging venue for value relevant news, and SA articles, in particular, provide information that accurately predicts a firm’s future earnings and future stock prices (Chen et al. 2014). However, while this evidence suggests that SA articles represent credible sources of information, no prior study has examined how personal financial incentives of social media participants (i.e., NPAs’ financial positions) affect contemporaneous reaction to these articles.

2.3 Hypotheses

We expect that NPA stock positions could increase investor response for at least three reasons. First, these NPAs likely conduct more diligent and careful research to form their opinions because they have a personal financial stake in the firm (i.e., they have “skin in the game”). Second, when expressing their opinions, NPAs with personal positions may withhold at least a portion of their private information, thus making the act of disclosing a stock position a signal in its own right, similar to professional analysts’ stock recommendations accompanying their detailed

⁹ As a rule, SA articles themselves do not uniformly include an author’s recommendation, though the author’s analysis may come with a recommendation to buy or sell a stock. In addition, SA uses keyword algorithms to generate categories of stock analysis, and two of those categories are “Long ideas” or “Short ideas.” However, not all articles by long (short) authors are tagged as “Long (Short) Ideas,” and not all authors writing a long and short idea have a position in the stock they recommend. Note that we use all SA articles (which have many different categories, including long and short ideas, among others) in our analyses.

analysis. Finally, investors with personal holdings have incentives to accelerate price formation in order to realize profits on their investment positions in a more timely fashion (e.g., Pasquariello and Wang 2018, Ljungqvist and Qian 2016).

As previously discussed, Chen et al. (2014) find that negative tone in a SA article is associated with lower returns over the following sixty trading days, implying that investors react to the information conveyed by SA articles *at the time* the article is published online. This is consistent with SA's own claim that "Seeking Alpha articles frequently move stocks" (Seeking Alpha 2016). Therefore, our first hypothesis tests whether an NPA's stock position is a signal about the NPA's beliefs regarding the valuation of the company (i.e., like an analyst's buy/sell recommendation or a manager's forecast) that is incremental to the information conveyed by the article. We expect the stock price to contemporaneously increase for the disclosure of long positions and to decrease for the disclosure of short positions, controlling for other information conveyed by the article (i.e., tone, length, rigor, numerical content, etc.).

H1: Holding constant other information conveyed by the article, investors respond to the disclosure of stock positions by NPAs.

If stock positions induce bias into SA articles, however, we do not expect to find support for H1.

Our next question is whether the credibility of NPA's voluntary disclosure is impaired or enhanced when (s)he has a financial position. As previously discussed, if the NPA has a financial position in the firm about which they write, this suggests that the NPA has "skin in the game" and, thus, might be more credible. If this is the case, we expect a *stronger* reaction to an NPA's tone when that NPA has a stock position (i.e., there is enhanced credibility) as compared to when the NPA takes no position. Our second hypothesis follows:

H2: Investors respond more strongly to tone in SA articles authored by NPAs with stock positions than by those with no stock positions.

On the other hand, investors could perceive that a stock position creates a conflict of interest. For example, NPAs could provide analysis that is intentionally biased in either positive or negative direction, and if it generates trading in the same direction, the NPA could personally profit from it. In addition, NPAs could report a position disclosure that they do not actually hold in order to move the market and profit from the movement. For example, they could report a short position that they do not have, expecting a short-term negative price response, and then purchase the stock at an artificially deflated price. In this case, the reaction to a NPA's tone should be *weaker*, as investors discount the tone of the information compared to when the NPA takes no position.

We test each hypothesis using short-window returns surrounding the release of each SA article, as we discuss in the next two sections. However, in additional analyses we also examine long-window returns (60 trading days) as in Chen et al. (2014) to assess whether any short-window effects persist or reverse. We discuss these results in Section 5.2.

3. Data and research design

3.1 Seeking Alpha data

We obtain news content from Seeking Alpha by systematically downloading all content published before July 7, 2015 (the date we performed the query). To ensure that we capture new analysis provided to the markets, we download “articles” (available at seekingalpha.com/article) rather than “news” (available at seekingalpha.com/news). The former represents long-form analysis whereas the latter represents shorter, news-flash-like content. Table 1 describes this beginning sample and sources of data-loss. In total, we obtain 487,197 SA articles.¹⁰ We then parse

¹⁰ SA offers a “pro” subscription that gives subscribers early access to content selected by editors. Per our discussion with the SA editor, during our sample period this access lasts 24 hours after which the article is made public for 30 days. After that 30-day period, the article is archived behind a paywall and available only to pro-subscribers. Due to our sampling procedures, we did download a limited number of “pro” articles that were published in the 30 days preceding our query of SA, but these articles were not accessible when we extracted comments at a later date and are therefore excluded from our final sample. Thus, our sample is fully comprised of articles which were available to the public on the timestamp appearing in the article. Note that as of mid-2018

each article to identify the article title, timestamp, referenced stocks (tickers), article content, authoring NPA, and position disclosure, each clearly delineated with specific HTML tags. The header information of the SA articles identifies tickers for referenced companies in two categories, “Primary” and “About.” Primary tickers are only identified when a company is the focus of the article and analysis, and the “About” tickers capture other mentioned companies. We exclude articles without a “Primary” ticker, as these articles often contain news summaries across the market or within a particular industry rather than substantive analysis regarding an individual firm. Excluding these summaries reduces our sample by 280,219 articles. Because our primary interest relates to the price effects of NPA positions, we delete another 58,378 articles with no position disclosure.¹¹

Disclosures generally, though not universally, follow the same basic format. At the beginning or end of each article, the NPA includes a statement such as “I am/we are long XXX,” “I/we have no positions in any stocks mentioned, and no plans to initiate any positions within the next 72 hours,” or “I am/we are short XXX.” However, in other instances positions are less clear, as the NPA may disclose complex option holdings or multiple positions in different stocks (i.e., long XXX and short YYY). Therefore, we use a two-stage procedure to code these disclosures.

First, we identify long positions by searching for the terms “long,” “hold,” or “own stock/shares.” We then capture the text following those words, stopping when a period or the word “may” or “short” is encountered. The latter two words indicate the beginning of a new position disclosure (i.e., “I am long ... and may...”). We repeat this procedure for possible short positions, looking for the word “short” and then capturing tickers until the word long or may, or a period.

this process appears to have changed. Immediate access to articles is limited to tickers included in a user’s portfolio, and access to archived content requires a pro subscription.

¹¹ Disclosures of positions were relatively rare until 2012 (no more than 20 percent each year). In more recent years, most (over 90 percent) of articles with a primary ticker designation include disclosures.

Note that for both long and short positions, we do not allow negating or qualifying words (no, not, none, neither, never, nobody, may, or plan) to occur within the five words preceding the position indicator. Finally, we search for cues that the NPA holds no position in any stocks. These include the terms “No Position,” “None,” or “May.”

Inspection of results suggests these search procedures are relatively accurate, but we do encounter complex disclosures that yield multiple classifications (i.e., long, short, and/or no position) or instances where we fail to identify any of the three positions. Further, disclosure of long or short positions could be in reference to stocks other than the stock about which the article is primarily written. Therefore, we further refine our disclosure coding as follows. First, in order to confirm a long or short position, we require the primary ticker of the article to match one of the tickers identified in the position disclosure. If the tickers do not match, we code the disclosure as “no position.” Second, inspection of disclosures where we fail to identify long, short, or no position cues suggests these are almost universally no-position disclosures, so we code them as such. Finally, we manually inspect 370 disclosures that our procedure tagged as both long and short. Based on this inspection, we code 80 of these articles as long, 44 as short. The remaining 246 disclosures correspond to unclear positions, usually involving both equity and option positions (i.e., own stock in X and short calls in X). We drop those from our sample, leaving 148,354 coded SA articles.¹²

We next attempt to match the primary tickers to the CRSP and Compustat header and history files. Approximately 4,700 tickers fail to match these header files. Manual inspection of the data suggests that the majority of these relate to ETFs or REITs, but we also fail to match a few large companies, such as Alphabet, Inc. and Under Armour, due to minor differences in tickers

¹² For brevity, we refer to articles authored by investors with long (short) positions as “long articles” (“short articles”), and to those holding no position as “no position articles”.

reported by SA and other data sources (e.g., GOOG vs. GOOGL). Therefore, we manually investigate each unmatched ticker that corresponds to at least 20 articles (approximately 200 stocks corresponding to 14 thousand articles) and identify a link to CRSP and Compustat where possible. In total, we lose 21,124 articles for stocks where we cannot identify a CRSP identifier (i.e., permno) upon which to merge. Finally, we lose another 11,358 articles that are missing any one of our basic control variables, and 10,920 with missing returns for any measurement window. This leaves us with a final sample of 104,952 articles. In most of our analyses, we collapse this dataset down to 86,741 unique firm-trading day combinations.¹³

3.2 Descriptive information on SA NPAs

Before moving to our research design and results, we first present some basic descriptive information about non-professional analysts (NPAs) that write on SA. We obtain this data from two sources. First, upon request, SA's Executive Editor provided us with basic demographic information they collect about the universe of SA NPAs. To supplement this data, we use a series of Python scripts to analyze the biographies posted on SA for the NPAs in our sample. We present the SA provided information (the information we generated) in Panel A (Panel B) of Table 2.

SA NPAs appear to largely consist of independent investors who are interested in establishing, building, and maintaining a reputation within the investment community. Based on SA's descriptive data, 75 percent of SA NPAs reveal their name and place of employment, suggesting that a majority of NPAs face reputational concerns not only online but also in their "day jobs." SA pays each NPA an average nominal wage of \$33.30 per month based on the number

¹³ While the sample attrition in Table 1 may at first glance appear dramatic, the majority of sample attrition is due to (1) the removal of SA articles that do not relate to one specific ticker symbol (i.e., that are industry or macroeconomy articles), (2) the requirement to disclose whether or not the NPA holds a position, and (3) the ability to match the identified ticker symbol with a firm listed in CRSP and Compustat. Thus, when put in context, our sample attrition is largely driven by factors that are necessary to answer our questions of interest. Given our research design, we do not believe these data restrictions induce any systematic biases in our sample.

of page-views. Thus, for most NPAs, monetary rewards do not appear to be the primary driver of producing high quality content. SA also reports that 27 percent of SA NPAs have their own independent investment blog, suggesting a substantial portion of NPAs invest significant time in the investment community beyond SA.¹⁴ Taken together, these results suggest that, on average, SA NPAs seek to establish a reputation within the financial community or to accelerate stock price formation rather than to earn money directly from their SA analysis activities.

Panel B provides the information we collected and coded for all NPAs in our sample as well as descriptive statistics by position (Short, No Position, Long). We first manually code each NPA as an individual, a company (a private investment firm, advisor, etc.), or anonymous. Similar to Panel A, approximately 27 percent of NPAs use an alias while the remaining 73 percent identify themselves. These statistics are similar across positions, except that NPAs we identify as companies more frequently disclose no positions. We also search for certain keywords in contributor biographies and find approximately 14 percent mention “Analyst,” 9 percent mention “MBA” (Masters degree in Business Administration), and 7 percent mention “CFA” (Certified Financial Analyst). These references are also similar across all positions, except that short-position NPAs appear twice as likely to have an MBA. Finally, SA reports “followers” for each contributor, much like Facebook or Twitter, and these followers are notified when contributors publish new content. In our sample, the average SA has a following of about 4,500 accounts, suggesting fairly wide dissemination of new content.

3.3 Empirical models

To test our hypotheses, we regress short-window abnormal returns on NPA position and a

¹⁴ SA prohibits authors from publishing the full content of their analysis in other locations. Therefore, it is unlikely articles written by authors with their own blog could be published in advance of clearing the editorial process at SA. Nonetheless, in a sensitivity analysis we exclude these observations from our sample. All of our results are quantitatively and qualitatively unchanged.

series of controls as presented below in [1] (i and t denote firm and time subscripts, respectively, and “[]” signifies a multi-day range):

$$\begin{aligned}
AbRet_{i,[t,t+1]} = & a_0 + a_1Long_{i,t} + a_2Short_{i,t} + a_3NegPct_{i,t} + a_4PosPct_{i,t} + a_5CogProc_{i,t} + \\
& a_6Numbers_{i,t} + a_7lWordCount_{i,t} + a_8ComNegPct_{i,[t,t+1]} + a_9ComPosPct_{i,[t,t+1]} \\
& + a_{10}DJPosPct_{i,t} + a_{11}DJNegPct_{i,t} + a_{12}IDJ_{i,t} + a_{13}Upgrades_{i,t} + \\
& a_{14}Downgrades_{i,t} + a_{15}ReviseUps_{i,t} + a_{16}ReviseDowns_{i,t} + a_{17}PosES_{i,t} + \\
& a_{18}NegES_{i,t} + a_{19}Guidance_{i,t} + a_{20}PosGuidance_{i,t} + a_{21}NegGuidance_{i,t} + \\
& a_{22}Edgar8K_{i,t} + a_{23}Volatility_{i,t} + a_{24}AbRet_{i,[t-60,t-3]} + a_{25}AbRet_{i,t-2} + \\
& a_{26}AbRet_{i,t-1} + a_{27}Size + a_{28}BTM_{i,t} + a_{29}InstOwn_{i,t} + a_{30}AnalystFollowers_{i,t-2} \\
& + a_{31}SAFollowers_{i,t-1} + \Sigma\gamma Industry_i + \Sigma\delta Year-Month + e_{i,t} \quad [1]
\end{aligned}$$

The dependent variable in [1] is the firm’s return measured over the two days starting on the day the article was published, adjusted by a matching size, market-to-book, and momentum portfolio return over the same period. If the article was published after-hours, on a weekend, or a holiday, we begin our return window on the first trading day following the article’s release.

In some cases, a stock has multiple articles written about it on the same day. If so, we follow Chen et al. (2014) and collapse the SA-derived data into firm-day observations to avoid including these firm-day combinations multiple times in our models. For instance, we compute *Long* and *Short* as the average number of articles on a given day that disclose long and short positions, respectively.¹⁵ To measure article tone, we count the number of positive and negative words, classified using word lists from Loughran and McDonald (2011), in all articles corresponding to a given trading day and divide each count by the total word count across articles, yielding *PosPct* and *NegPct*, respectively.¹⁶

Based on H1, we expect a positive (negative) coefficient on *Long* (*Short*). To test H2, we estimate [1] separately for four different cross-sections, which we denote with *Position*. *Position*

¹⁵ We provide detailed variable definitions in Appendix A.

¹⁶ Following Loughran and McDonald (2011), we do not code words as positive if they are preceded by a negating word (no, not, none, neither, never, or nobody).

equals 0 if both *Long* and *Short* equal 0 on a given trading day. *Position* equals -1 (1) if *Short* exceeds *Long* (*Long* exceeds *Short*) on day t . If *Short* and *Long* are equal and non-zero, we exclude these days from our sample. H2 suggests that investors respond more strongly to tone expressed by NPAs with “skin in the game.” Specifically, H2 predicts that the coefficient on both *NegPct* and *PosPct* exhibits stronger significance (in the expected direction) in partitions where NPAs hold a financial position (*Position* is non-zero).

Our control variables attempt to isolate other news that may affect current period returns and SA article content.¹⁷ First, we control for the volume and rigor of analysis, as an NPAs’ personal financial position might simply reflect the depth of their analysis. Specifically, we include the proportion of words reflecting higher cognitive effort from James Pennebaker’s LIWC package. Per Pennebaker and Francis (1996), words such as “believe,” “cause” and “consider” reflect a higher degree of cognitive engagement with written material. We also include the number of numbers in the article, scaled by article length, to capture the degree of specificity of the analysis (*NumCount*), and the natural log of the total number of words in the article (*lWordCount*) in case there is asymmetry in how investors respond to length. Second, we control for other “news” appearing alongside the SA article. Chen et al. (2014) find that comments following SA articles provide value relevant information. Therefore, we separately download comments for each article in our sample and code positive and negative linguistic tone for comments (*ComPosPct* and *ComNegPct*) using the same procedure as *PosPct* and *NegPct*. We restrict our comment sample to those disclosed between the date of the article and the second trading day in our return window. We also control for the tone of the business press using news disseminated by the Dow Jones

¹⁷ The long-form nature of SA articles makes it unlikely that an event on day t leads to an article written on day t . Furthermore, discussions with an editor at SA suggest that the editorial process can be lengthy—as long as 12 hours in some cases. This delay makes it unlikely that reactions to SA content reflect contemporaneously issued news. Nonetheless, we control for several aspects of contemporaneous news in our models, and we conduct additional analyses in Section 5 to rule out the alternative explanation that contemporaneous events explain our findings.

newswire (*DJPosPct* and *DJNegPct*) as well as an indicator, *IDJ*, equal to 0 on days where there is no Dow Jones content (*DJPosPct* and *DJNegPct* set to 0 on these days).

We also control for the presence of several other significant news events in the four-day window ending on the day of the article's publication. Specifically, we control for analyst upgrades, downgrades and forecast revisions (*Upgrades*, *Downgrades*, *ReviseUps* and *ReviseDowns*) and positive and negative earnings surprises (*PosES* and *NegES*). Furthermore, we control for the presence or absence of management guidance (*Guidance*) as well as its sign (*PosGuidance* and *NegGuidance*) and for the presence or absence of an 8-K filing. Finally, we control for several return-based measures of news and uncertainty. Specifically, we compute pre-disclosure volatility (*Volatility*), which captures uncertainty in the calendar month preceding the article's release, and pre-article stock performance over three separate windows (day $t-60$ to $t-3$, day $t-2$, and day $t-1$).

In addition to these determinants, we include controls for firm size (*Size*), growth (*BTM*), institutional ownership (*InstOwn*), and following by both professional analysts and SA readers (*AnalystFollowers* and *SAFollowers*).¹⁸ Finally, all models include “year-month” and industry fixed effects.

3.4 Descriptive Statistics

Table 3 presents descriptive statistics for our 86,741 firm-day observations. Variables marked with “*” are scaled by 100 to facilitate presentation of descriptive statistics. Means and medians for each return metric (*AbRet*) all hover about 0, suggesting fairly symmetric return distributions. Statistics for *Short* (*Long*) suggest that approximately 2 (27) percent of articles are

¹⁸ We include *SAFollowers* to control for the author's reputation and/or ability. Our results are unchanged if, in addition to *SAFollowers*, we also include as control variables (1) a measure of the number of prior posts, (2) the number of prior posts about the target firm, (3) the author “track record” following Chen et al., and (4) whether or not the author is deemed a “financial professional” (coded as 1 if the author mentions “analyst” or “CFA” in his or her bio or if the account name appears to be a business).

authored by NPAs with a short (long) position. Our tone measures (*PosPct* and *NegPct*) suggest that NPAs use only about 1.3 (1.5) percent of negative (positive) words in articles. Chen et al. (2014) report similar statistics for negative words (they do not report statistics for positive words). We observe similar statistics for the language used in the comments section to the articles. On average, NPAs appear to exert fairly significant cognitive effort (mean *CogProc* of 9.1 percent) and include a fair amount of numerical content (i.e., numbers account for about 4.5 percent of total words). We also find that 10 (15) percent of articles co-occur with an analyst upgrade (downgrade). Forecast revisions occur in approximately 25 percent of our sample and frequently come in clusters (as evidenced by means greater than 1). Furthermore, 9 (5) percent our sample corresponds to the period following a positive (negative) earnings surprise, 7 percent of articles also occur near management guidance.¹⁹ Finally, 57 percent of SA-article days have at least one other co-occurring news-item as tracked by the Dow Jones newswire (*IDJ*), and the use of tone-words in these articles is relatively sparse (less than one percent for both positive and negative words).

Table 4 presents correlations among our variables. Bolded correlations are significantly different from zero ($p < 0.05$). Consistent with H1, we observe a significantly positive (negative) correlation of 0.06 (-0.06) between *Long* (*Short*) and $AbRet_{i,[t,t+1]}$. We also observe positive (negative) correlations between $Abret_{i,[t,t+1]}$ and both *PosPct* and *ComPosPct* (*NegPct* and *ComNegPct*), implying market movement in the direction consistent with the tone of SA articles. The tone of Dow Jones content exhibits weaker correlations to short-window returns (-0.03 and 0.01 for negative and positive tone, respectively). Interestingly, few non-SA related variables relate significantly to *Short* and *Long*. We observe positive correlations between *Volatility* and

¹⁹ *PosGuidance* and *NegGuidance* take the value of 1 if management issues a forecast that is greater or less than the analyst consensus before the guidance is issued, respectively, while *Guidance* is equal to 1 if any forecast is issued. The sum of the mean values for *PosGuidance* and *NegGuidance* slightly exceeds the value of *Guidance* because managers may issue multiple forecasts of varying horizons in a given window with different news (i.e., *PosGuidance* and *NegGuidance* could both equal 1)

both *Long* (0.04) and *Short* (0.09), implying that NPAs with positions more likely publish content when uncertainty is relatively high. Long and short positions also appear more likely for smaller firms (-0.09 and -0.07) with lower institutional ownership (-0.10 and -0.11) and analyst following (-0.05 and -0.09), suggesting NPAs with positions target stocks in relatively poorer information environments. We also observe a correlation of 0.03 between the abnormal return over the prior quarter ($AbRet_{i,[t-60,t-3]}$) and *Short*, suggesting that past news plays, at most, a minor role in these NPAs' decisions to publish. With respect to SA-article derived variables, we find a greater (smaller) intensity of negative (positive) words for short NPAs, suggesting that these NPAs write content consistent with their positions. Interestingly, we find that *Long* NPAs use fewer negative words, though not necessarily more positive words. The tone of comments tends to follow the tone of articles, and *Short* (*Long*) NPAs tend to incite more negative and less positive (more negative and more positive) comment sentiment. We also find that long and short NPAs write articles indicating greater cognitive effort but include fewer numbers per word of text.

4. Empirical Results

4.1 Test of H1

H1 predicts that there is an investor reaction to position disclosures in SA articles at the time they are published, or that the disclosure of a long position (*Long*) generates a positive abnormal return in the short-window surrounding the articles release and disclosure of a short position (*Short*) generates a negative abnormal return. We report results for these predictions in Panels A and B of Table 5 using equation [1].²⁰

²⁰ In all tables, we multiply the dependent variable by 100 to facilitate presentation of coefficient estimates. All specifications also include industry and year-month fixed effects. Significance is assessed from standard errors clustered by year-month to correct for cross-sectional correlation in returns.

Panel A presents results using the 86,741 firm-day combinations in our sample. Columns 1 through 3 present results using various sets of control variables. We begin with a baseline model in column 1 that includes variables measured from SA articles themselves: *PosPct*, *NegPct*, *Long*, *Short*, *CogProc*, *Numbers*, *lWordCount*, *ComPosPct*, and *ComNegPct*. As presented, we find strong support for H1 as the coefficient on *Long* is significantly positive (0.431, t -statistic = 11.92) and the coefficient on *Short* is significantly negative (-1.045, t -statistic = -9.12). These coefficients imply a 2-day abnormal return of 0.4 (-1.0) percent attributable to the disclosure of a long (short) position. The economic magnitude of our results are in line with Chen et al. (2014), particularly in light of the fact that we are examining the immediate reaction to the news and not the subsequent drift. We also observe significant coefficients on both *PosPct* and *NegPct* (12.151, t -statistic = 6.27 and -16.43, t -statistic = -8.52, respectively). A one standard-deviation increase in *PosPct* (*NegPct*) corresponds to a return of 0.1 (-0.1) percent, which is not insignificant in a 2-day window but far smaller than the return attributable to non-professional analysts' (NPAs') positions.

We next introduce controls for non-Dow Jones related news content (column 2), such as analyst revisions and management guidance, and DJ content (column 3) and continue to find strong support for H1. Coefficients on *Long* all hover around 0.4 (t -statistics > 10.0) while coefficients on *Short* suggest incremental returns between 1.05 and 1.16 percent (t -statistics < -9.12). Coefficients on *NegPct* and *PosPct* also exhibit similar magnitudes to column 1. We next consider various subsamples that reduce the likelihood articles appear alongside contemporaneous news. Specifically, in columns 4 and 5, we re-estimate our full model (corresponding to column 3) after removing observations with Dow Jones content ($IDJ = 1$) and after removing observations with DJ content and earnings surprises ($IDJ = 1$ or $PosES = 1$ or $NegES = 1$). All results continue to hold and the coefficients on *Long* and *Short* exhibit noticeably larger magnitudes compared to the

column 3 estimates, suggesting a greater price response to SA content on days without concurrent news events. Column 6 shows results after excluding articles occurring in the three days preceding an analyst recommendation, forecast revision, earnings announcement, or management guidance, and column 7 only retains days on which a single article is published (since days with multiple articles likely correspond to significant firm events). Our results are similar in both cases.²¹ Overall, the results presented in Table 5 provide strong support for H1, suggesting that stock positions convey information about the NPA's overall opinion of the firm and that investors perceive NPAs to be credible.

One immediate concern related to these results is that NPA positions correlate with some unobserved NPA characteristic, such as reputation or writing style, which drives the results found in support for H1 and for which we have been unable to control. To address this possibility, we supplement equation [1] with NPA fixed effects and estimate this model at the article rather than firm-day level. Panel B reports results of this test. For brevity, we only report coefficients on SA-related variables. As presented, we continue to find strong support for H1, as both *Long* and *Short* exhibit highly significant associations (*t*-statistics between 4.43 and 7.29 in magnitude) with returns in the 2-day window following the article's publication.²² The coefficients in Panel B of Table 5 suggest a positive 2-day return of approximately 0.3 percent for long positions, and disclosure of a short position corresponds to negative returns between 0.60 and 1.0 percent over the same period. These effects are similar, though smaller than those presented in Panel A.

4.2 Test of H2

²¹ Our results also hold if we exclude articles published on days with *any* information events in the surrounding period.

²² To the extent that the absolute (rather than signed) reaction to an author's work varies with his or her reputation, using signed returns in Panel B may not adequately control for unobserved author characteristics. Therefore, in untabulated tests, we replace the dependent variable with the natural log of 1 plus the absolute value of $AbRet_{i,[0,1]}$ and repeat tests in Panel B of Table 5. While we continue to find significant coefficients on *Short* in all specifications (*t*-statistics between 2.54 and 4.75), coefficients on *Long* are significant only in full-sample models.

H2 predicts that investors respond more strongly to the tone (*PosPct* and *NegPct*) of articles authored by those with stock positions compared to those with no position. We report results related to H2 in Table 6, where we use *Position* as a partitioning variable and estimate equation [1] using four separate subsamples: (1) $Position = 0$, (2) $|Position| = 1$, (3) $Position = -1$, and (4) $Position = 1$.²³ The final four columns of Table 6 report tests of differences of coefficients across the partitions (i.e., the header “1-2” reports the significance of the difference in coefficients between columns 1 and 2).

Columns 1 and 2 present our formal test of H2. Specifically, H2 suggests that the coefficients on *NegPct* and *PosPct* should be stronger (larger in magnitude) in column 2 than in column 1. The magnitude of the coefficient on *NegPct* in column 2 (-31.925) exceeds that in column 1 (-11.696), and this difference is highly significant ($p < 0.01$). We observe a similar pattern for *PosPct*. The column 2 coefficient (23.3) significantly exceeds the column 1 value (8.862, difference significant at $p < 0.01$). While we make no predictions of how coefficient patterns differ over the remaining columns in Table 6, we present the remaining tests-of-differences for completeness. Comparing coefficient estimates in column 3 (4) to column 1 provides an indication of whether positive and negative tone by short (long) NPAs is considered more credible than NPAs with no position. Interestingly, the magnitudes of coefficients on *PosPct* in column 3 and *NegPct* in column 4 are larger than the same coefficients in column 1, though these differences are only marginally significant (one-tailed p -values of 0.06 and 0.04, respectively). Thus, positive (negative) tone in articles written by NPAs is more credible when they hold short (long) positions, suggesting that an NPA is most credible when reporting

²³ We present results for H2 using sample partitions. To assess significance, we estimate equation [1] using a series of fully interacted models. Reported significance levels reflect the significance of the interaction distinguishing the two compared columns. We report one-tailed p -values when comparing “position” articles vs. no position (i.e., columns 2, 3, or 4 vs. column 1) and two-tailed otherwise.

information contrary to their position. Comparing column 3 (short positions) to column 4 (long positions) suggests no significant differences.

In sum, we find strong support for H2. Investors appear to perceive tone by NPAs holding positions to be more credible than those with no position. Additional evidence suggests that these results are primarily driven by the response to tone that is contrary with an author's position (i.e., positive tone for short authors and negative tone for long authors). Importantly, though, we find no evidence that even position-consistent tone is discounted, implying investors perceive no bias related to NPAs having "skin in the game."

5. Additional analysis

5.1 Alternative explanations: Contemporaneous events leading to observed stock price reactions

One possible concern with our results is that our variables of interest capture some other article attribute or contemporaneous event. We believe this to be unlikely because the long-form of SA articles, which frequently include extensive tables, charts, and links to detailed analysis, makes it unlikely that a non-professional analyst (NPA) observes an event and immediately produces such an article. In addition, SA articles undergo an editorial process which, per our discussions with an SA executive, averages 4.5 hours and can take up to 12. Finally, with respect to H2, contemporaneous events would need to correlate with not only the content of SA articles, but also the position of the NPA writing the article, and an inspection of SA articles reveals no systematic differences in articles authored by positioned NPAs (*Short* or *Long* equaling 1) compared to no-position NPAs. Nevertheless, in this section we perform several additional tests designed to mitigate the likelihood that our results are due to contemporaneous firm economic events.

5.1.1 Controlling for firm news around the SA article release date and time

As just discussed, one alternative explanation for our results is that the stock returns we observe are not a reaction to the SA article but are instead a reaction to contemporaneous firm news events. In our main tests, we attempt to rule out this explanation by including a host of control variables related to contemporaneous news, including the tone of business press articles, analyst revisions and recommendations, and management guidance. To further mitigate this issue, we take advantage of the fact that the SA editorial process takes an average of 4.5 hours to complete and assume that any article published on SA before 1 pm (i.e., the first 3.5 trading hours of the day) must have been submitted to SA before the market opened for the day. We then redefine the day 0 return as $(\text{closing price} - \text{opening price}) / \text{opening price}$ (all measured on day 0). Therefore, any overnight or pre-market news impounded into the opening price is excluded from our returns.

Using this revised measure of returns and limiting the sample to articles posted in the first few hours of trading, we repeat our analyses from Table 5 and present the results in Table 7. As shown, we continue to find significant coefficients on both *Long* and *Short*, and the economic significance of these effects is relatively unchanged. These results add further assurance that our stock price results are not explained by contemporaneous news events and are, instead, a reaction to the SA articles themselves.

5.1.2 Interactions between rigor of analysis and position

While we contend that disclosure of stock positions provides a value relevant signal to market participants, we recognize that simply saying “I am long...” without any other support would likely garner little investor reaction. Therefore, we expect that the reaction to NPA positions increases with the amount of information and quality of analysis presented alongside their position disclosure. To test this conjecture, we use cognitive effort (*CogProc*), the number of numbers (*Numbers*), and article length (*IWordCount*) as proxies for the quality and volume of information

in the article and interact these variables with both *Long* and *Short* in equation [1]. We expect each interaction to load in the same direction as the position (e.g., positive for *Long*lWordCount* and negative for *Short*lWordCount*).²⁴

Results from these analyses are presented in Table 8. We include the same subsamples as in earlier analyses (all observations, excluding DJ, excluding DJ and earnings surprises, excluding post-article information events). In all four columns, coefficients on *Long* and *Short* remain significant at less than the 1% level, and magnitudes are similar to those shown in Table 5. Most importantly, we observe several significant coefficients on interactions in the expected direction. That is, in all specifications, the interaction between *Long* and *lWordCount* is significantly positive (*t*-statistics between 2.84 and 3.66) and the interaction between *Short* and *lWordCount* is significantly negative (*t*-statistics between -3.69 and -5.85). We also find some evidence that the number of numbers (*Numbers*) increases the credibility of short positions (*t*-statistics of -2.18, -1.72 and -1.86 in columns 1, 2 and 3, respectively). Thus, investor reaction to NPA positions appears to increase with article length, as expected, and, to a lesser extent, numerical information enhances the reaction to short position disclosures. More importantly, the fact that the main effects for *Short* and *Long* continue to hold provides further evidence that that NPA positions themselves convey meaningful information to readers and are not capturing only dimensions of article quality.

5.1.3 First-time vs. Repeated Disclosures

To further support that NPA position, and not an unidentified correlated omitted variable, explains the reaction to SA content, we next examine whether the reaction to position is stronger in the NPA's first article publishing his or her stock position. Specifically, we sort our sample of

²⁴ We center these three variables about 0 for this analysis to maintain interpretability of coefficients on *Long* and *Short*. In other words, main effects on *Long* and *Short* represent the response to position for average levels of the interacted variables. This is especially useful, as another explanation for our results related to H1 is that, for articles coded by NPAs with a long (short) position, article length captures positive (negative) tone not measured in *PosPct* (*NegPct*). If this were the case, then we would observe a significant interaction between each position variable and *lWordCount* but insignificant main effects on *Long* and *Short*.

SA content by firm (i.e., primary ticker), NPA, and date, and identify whether the article marks the first time an NPA discloses a position about the subject firm. We denote this article using an indicator variable, *FirstDisc*. We then estimate equation [1], including interactions between *FirstDisc* and both *Long* and *Short*. We expect that the relations between returns and both *Short* and *Long* are stronger the first time an NPA discloses a position.²⁵ In addition, we include interactions between *FirstDisc* and other article attributes (*CogProc*, *Numbers*, *lWordCount*, *NegPct*, *PosPct*, *ComNegPct*, and *ComPosPct*) because relations between those variables and returns may vary depending on how often the NPA writes about a given firm. However, we make no predictions related to these interactions.

Table 9 presents results from this analysis. For brevity, we only include coefficients on the interactions between SA-related variables and *FirstDisc* and suppress tabulation of other coefficients. We include the same subsamples used throughout the paper. Consistent with expectations, we observe a highly significant, negative coefficient on the interaction between *Signal* and *Short* for first disclosures (*t*-statistics between -4.5 and -5.8) and a significantly positive coefficient on the interaction between *Signal* and *Long* (*t*-statistics between 3.11 and 3.43).²⁶

5.2 SA content and long-run returns

As discussed, Chen et al. (2014) document a significantly negative return between the percentage of negative words in SA articles and returns over the subsequent quarter (approximately 60 trading days). In our main tests, we focus on contemporaneous pricing of SA content and how position affects this pricing under the assumption that markets are generally efficient. We now repeat these analyses (Tables 5 and 6) replacing our short window return ($AbRet_{i,[t,t+1]}$) with the

²⁵ One may argue that the first-time disclosure of a position should be the only time this knowledge matters. However, multiple articles disclosing the same position affirm the author's beliefs over time, thus providing additional relevant signals.

²⁶ We also considered whether a change in disclosures represents a credible signal. However, this only occurs in about 8 percent of articles.

post-event return over the following 60 days ($AbRet_{i,[t+3,t+60]}$), as in Chen et al. (2014). We also examine returns during the window three to five days after the call, three to ten days and three to twenty days.

Table 10 replicates Table 5 using post-event returns. We use all controls from [1] and add $AbRet_{i,[t,t+1]}$ and tone from comments over the duration of the full return window. We include the same models as in Table 5. Unlike Chen et al. (2014), we do not observe associations between $NegPct$ and $PosPct$ and post-event returns.²⁷ We do, however, find a significant, negative association between 60-day returns and $Short$, suggesting that the market does not fully impound the information content of a short position.

6. Conclusion

Motivated by concerns that financial positions present a conflict of interest that impairs an analyst's objectivity, we examine investor perceptions of the financial positions of non-professional analysts (NPAs) providing stock analysis on the social media outlet SeekingAlpha (SA) and offer two primary findings. First, NPA positions contribute directly to short-window returns surrounding the article's publication, holding constant the information in the article (i.e., tone, length, rigor, numerical content, etc.) as well as contemporaneously issued news (i.e., from professional analysts and the business press). Economically, long NPA authored articles correspond to a 2-day return of 0.4 percent, while short NPA authored articles correspond to returns of -1.0 percent over the same period. These findings suggest that an NPA's stock positions convey credible information to investors. Second, we find that the price response attributed to article tone is significantly stronger for articles authored by NPAs with stock positions, and these

²⁷ In untabulated tests, we replicate their result using size and book-to-market matched portfolio returns as well as simple market-adjusted returns. Thus, the difference in our result appears to be driven by return definition (ours corrects for momentum, theirs does not) rather than differences in sample period or control variables. Additionally, Chen et al. (2014) still provide important evidence that SA articles predict future earnings news.

effects appear driven mostly by tone contrary to an author's position. Overall, our results suggest that the disclosure of an NPA's financial positions enhances their credibility with investors.

Three additional analyses bolster our findings. First, we show that limiting our sample to a subset of articles that are unlikely to be affected by contemporaneous news has no effect on our results. Second, we suspect that position disclosures are more informative when the NPA provides more information relevant to the firm, increasing his or her credibility. Consistent with this conjecture, we show that the association between NPA position disclosures strengthen with article length, rigor, and numerical content. Thus, position disclosures matter more when accompanied by more analysis. Finally, we suspect that first-time disclosures of positions garner greater responses because the information is "new." Consistent with this expectation, we show that reaction to both short and long disclosures is significantly stronger the first time an NPA discloses a position.

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APPENDIX A: VARIABLE DEFINITIONS

Variable	Definition
<i>AbRet_{i,t}</i>	The firm's return measured on day <i>t</i> or over days [<i>t</i> to <i>t+k</i>] adjusted by a matching size, market-to-book and momentum portfolio return over the same period. If the article was published after-hours, on a weekend, or holiday, day <i>t</i> equals the first trading day following the article's release. (winsorized)
<i>Position</i>	Takes value -1 if <i>Short</i> exceeds <i>Long</i> on a given day, 1 if <i>Long</i> exceeds <i>Short</i> on a given day, and 0 if <i>Long</i> and <i>Short</i> both equal 0. On days where <i>Long</i> = <i>Short</i> and both <i>Long</i> and <i>Short</i> are non-zero, <i>Position</i> is undefined.
<i>Short</i>	The percentage of articles about firm <i>i</i> on day <i>t</i> in which the NPA discloses a short position.
<i>Long</i>	The percentage of articles about firm <i>i</i> on day <i>t</i> in which the NPA discloses a long position.
<i>NegPct</i>	The percentage of non-negated words for all SA articles on day <i>t</i> that are classified as having negative sentiment using Loughran and McDonald's (2011) dictionary. (winsorized)
<i>PosPct</i>	The percentage of non-negated words for all SA articles on day <i>t</i> that are classified as having positive sentiment using Loughran and McDonald's (2011) dictionary. (winsorized)
<i>CogProc</i>	Count of cognitive processing words, such as "believe," "cause" and "consider" from LIWC, a commonly used psycholinguistic software package.
<i>Numbers</i>	The number of numbers, either as strings of digits and valid punctuation or written out in letters, divided by the total number of words appearing in SA articles about a firm on a given day. (winsorized)
<i>lWordCount</i>	The natural log of the total number of words appearing in SA articles about a firm on a given day. (winsorized)
<i>ComNegPct_{i,t}</i>	The percentage of non-negated words appearing in comments posted between day <i>t</i> and <i>t+k</i> about the SA article classified as having negative sentiment using Loughran and McDonald's (2011) dictionary.
<i>ComPosPct_{i,t}</i>	The percentage of non-negated words appearing in comments posted between day <i>t</i> and <i>t+k</i> about the SA article classified as having positive sentiment using Loughran and McDonald's (2011) dictionary.
<i>DJNegPct</i>	The percentage of non-negated words in all Dow Jones news content published on day <i>t</i> , or in the days between article publication and first trading day if different, classified as having positive sentiment using Loughran and McDonald's (2011) dictionary. (winsorized)
<i>DJPosPct</i>	The percentage of non-negated words in all Dow Jones news content published on day <i>t</i> , or in the days between article publication and first trading day if different, classified as having negative sentiment using Loughran and McDonald's (2011) dictionary. (winsorized)
<i>IDJ</i>	An indicator equaling 1 if there is no Dow Jones content about the firm published on day <i>t</i> , or in the days between article publication and first trading day if different.
<i>Upgrades</i>	The number of analysts revising recommendations upward on day <i>t</i> , or in the days between article publication and first trading day if different. (winsorized)
<i>Downgrades</i>	The number of analysts revising recommendations downward on day <i>t</i> , or in the days between article publication and first trading day if different. (winsorized)

<i>ReviseUps</i>	The number of analysts issuing earnings forecasts exceeding the prevailing consensus between day $t-3$ and the later of the article's publication date or first trading day following the article's publication date.
<i>ReviseDowns</i>	The number of analysts issuing earnings forecasts lower than the prevailing consensus between day $t-3$ and the later of the article's publication date or first trading day following the article's publication date.
<i>PosES</i>	Indicator equaling 1 if the firm announces earnings exceeding the most recent consensus estimate according to IBES on day t , or in the days between article publication and first trading day if different. (winsorized)
<i>NegES</i>	Indicator equaling 1 if the firm announces earnings below the most recent consensus estimate according to IBES on day t , or in the days between article publication and first trading day if different. (winsorized)
<i>Guidance</i>	An indicator variable equaling 1 if the firm issues at least one piece of earnings guidance between day $t-3$ and the later of the article's publication date or first trading day following the article's publication date.
<i>PosGuidance</i>	An indicator variable equaling 1 if the firm issues at least one piece of earnings guidance between day $t-3$ and the later of the article's publication date or first trading day following the article's publication date that exceeds the prevailing analyst consensus on the forecast date.
<i>NegGuidance</i>	An indicator variable equaling 1 if the firm issues at least one piece of earnings guidance between day $t-3$ and the later of the article's publication date or first trading day following the article's publication date that falls below the prevailing analyst consensus on the forecast date.
<i>Volatility</i>	The sum of squared daily returns in the calendar month preceding day t . (winsorized)
<i>Edgar8K</i>	An indicator variable equaling 1 if the firm issues at least one 8-K filing between day $t-3$ and the later of the article's publication date or first trading day following the article's publication date.
<i>Size</i>	The natural log of the market value equity as of the end of the month prior to the article's publication date. (winsorized)
<i>BTM</i>	The book value of equity as of the end of the most recent fiscal year divided by the market value of equity as of the end of the prior year. (winsorized)
<i>InstOwn</i>	The proportion of shares owned by institution investors per the quarterly Thomson Reuters ownership report closest but prior to the article publication date. (winsorized)
<i>AnalystFollowers</i>	The natural log of 1 plus the number of analysts issuing estimates in the IBES summary report in the month prior to the article's publication. (winsorized)
<i>SAFollowers</i>	The natural log of 1 plus the number of followers reported on the author's bio page as of the date the biographies were harvested. (winsorized)
<i>FirstDisc</i>	Indicator taking value of 1 the first time an NPA discloses a given position about a firm.

TABLE 1: SAMPLE ATTRITION

Seeking Alpha Articles Downloaded as of July 7, 2015	487,197
Articles missing primary ticker designation	(280,219)
Articles missing position disclosure	(58,378)
Articles with ambiguous position disclosure	(246)
Articles with Successfully Coded Disclosures	<hr/> 148,354
Articles not linked to CRSP	(21,124)
Articles missing other controls	(11,358)
Articles missing returns for any period	(10,920)
Total Articles in Sample	<hr/> 104,952
Unique Firm-day combinations	<hr/> 86,741 <hr/>

TABLE 2: NPA CHARACTERISTICS

PANEL A: NPA CHARACTERISTICS AS DESCRIBED BY SEEKINGALPHA

	# of NPAs	% of Total
Total NPAs	13,680	
NPAs with independent blogs	3,711	27.13%
Anonymous NPAs	3,358	24.55%
Financial Professionals	4,891	35.75%
Students (Young Investors)	1,056	7.72%
Company Executives/C-Level	748	5.47%
Monthly Average Payment to NPAs in 2016	\$33.30	

Panel B: NPA characteristics harvested from biographies

Characteristic	Total	Position = - 1	Position = 0	Position = 1
Individual NPA	56.16%	60.51%	53.03%	63.53%
Company NPA	16.37%	8.89%	19.27%	9.88%
AnonymousNPA (alias)	27.47%	30.60%	27.70%	26.59%
	100.00%	100.00%	100.00%	100.00%
Includes "Analyst" in bio	14.28%	16.11%	14.51%	13.52%
References blog or website (other than LinkedIn)	42.06%	41.04%	42.07%	42.12%
Mentions "MBA" in bio	8.76%	16.67%	8.74%	8.03%
Mentions "CFA" in bio	6.62%	7.00%	7.26%	4.97%
Followers at time bio page was downloaded (mean)	4,486	3,850	4,487	4,547

Table 2 reports descriptive statistics for NPA characteristics. Panel A reports information provided by SeekingAlpha and Panel B reports information pertaining to articles in our sample. For Panel B, we downloaded each NPA's bio from SeekingAlpha (<http://seekingalpha.com/author/...>) and used hand-coding or textual analysis to harvest select information. We manually coded each NPA as an individual, a company, or an alias. Remaining information was systematically extracted from each bio page.

Table 3: Descriptive Statistics

Variable	n	Mean	Std. Dev	25%	50%	75%	Position = -1		Position = 0		Position = 1		Tests of Differences		
							Mean	50%	Mean	50%	Mean	50%	-1 vs. 0	-1 vs. 1	0 vs. 1
<i>AbRet_{i,[t,t+1]}*</i>	86,741	0.11	3.62	-1.26	0.02	1.35	-1.11	-0.61	0.03	-0.02	0.39	0.12	0.00	0.00	0.00
<i>AbRet_{i,[t+3,t+60]}*</i>	86,641	-0.56	16.45	-8.77	-0.63	7.34	-1.79	-1.75	-0.47	-0.58	-0.66	-0.64	0.00	0.01	0.12
<i>AbRet_{i,[t-2]}*</i>	86,741	0.02	2.39	-0.88	-0.02	0.86	-0.10	-0.14	0.02	-0.01	0.02	-0.01	0.06	0.08	0.65
<i>AbRet_{i,[t-1]}*</i>	86,741	0.05	2.79	-0.89	-0.01	0.91	-0.10	-0.12	0.04	-0.01	0.08	0.00	0.08	0.02	0.05
<i>AbRet_{i,[t-60,t-3]}*</i>	86,741	0.30	18.37	-8.87	-0.53	7.78	3.23	0.44	0.17	-0.43	0.32	-0.83	0.00	0.00	0.29
<i>Position</i>	86,741	0.27	0.51	0.00	0.00	1.00	-1.00	-1.00	0.00	0.00	1.00	1.00	.	.	.
<i>Short</i>	86,741	0.02	0.15	0.00	0.00	0.00	0.88	1.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00
<i>Long</i>	86,741	0.27	0.43	0.00	0.00	0.67	0.00	0.00	0.00	0.00	0.90	1.00	0.00	0.00	0.00
<i>NegPct*</i>	86,741	1.33	0.87	0.69	1.18	1.79	2.09	1.93	1.32	1.16	1.28	1.16	0.00	0.00	0.00
<i>PosPct*</i>	86,741	1.47	0.75	0.94	1.38	1.90	1.03	0.96	1.49	1.40	1.46	1.39	0.00	0.00	0.00
<i>CogProc*</i>	86,741	9.10	3.35	7.30	9.41	11.33	10.86	11.41	8.78	9.08	9.64	9.90	0.00	0.00	0.00
<i>Numbers*</i>	86,741	4.54	3.17	2.37	3.90	5.89	4.12	3.53	4.59	3.91	4.45	3.92	0.00	0.00	0.00
<i>lWordCount</i>	86,741	6.76	0.63	6.36	6.74	7.13	7.07	7.08	6.67	6.69	6.93	6.88	0.00	0.00	0.00
<i>ComPosPct_{i,[t,t+1]}*</i>	86,741	1.02	1.19	0.00	0.91	1.53	0.90	0.92	0.97	0.73	1.16	1.13	0.00	0.00	0.00
<i>ComNegPct_{i,[t,t+1]}*</i>	86,741	1.00	1.08	0.00	0.89	1.63	1.49	1.55	0.93	0.64	1.09	1.12	0.00	0.00	0.00
<i>ComPosPct_{i,[t+3,t+60]}*</i>	86,741	0.59	1.19	0.00	0.00	0.87	0.71	0.47	0.52	0.00	0.74	0.00	0.00	0.15	0.00
<i>ComNegPct_{i,[t+3,t+60]}*</i>	86,741	0.59	1.11	0.00	0.00	0.96	1.18	0.99	0.50	0.00	0.74	0.00	0.00	0.00	0.00
<i>DJPosPct*</i>	86,741	0.67	0.99	0.00	0.00	1.24	0.72	0.00	0.66	0.00	0.68	0.00	0.01	0.08	0.01
<i>DJNegPct*</i>	86,741	0.51	0.71	0.00	0.00	0.96	0.52	0.00	0.51	0.00	0.51	0.00	0.36	0.57	0.37
<i>IDJ</i>	86,741	0.57	0.49	0.00	1.00	1.00	0.56	1.00	0.56	1.00	0.58	1.00	0.85	0.19	0.00
<i>Upgrades</i>	86,741	0.10	0.34	0.00	0.00	0.00	0.12	0.00	0.10	0.00	0.09	0.00	0.02	0.01	0.55
<i>Downgrades</i>	86,741	0.15	0.43	0.00	0.00	0.00	0.16	0.00	0.14	0.00	0.15	0.00	0.05	0.12	0.24
<i>ReviseUps</i>	86,741	1.54	3.91	0.00	0.00	1.00	1.27	0.00	1.52	0.00	1.62	0.00	0.00	0.00	0.00
<i>ReviseDowns</i>	86,741	1.74	4.09	0.00	0.00	2.00	1.76	0.00	1.70	0.00	1.81	0.00	0.58	0.60	0.00
<i>PosES</i>	86,741	0.09	0.29	0.00	0.00	0.00	0.09	0.00	0.10	0.00	0.09	0.00	0.49	0.62	0.00
<i>NegES</i>	86,741	0.05	0.21	0.00	0.00	0.00	0.05	0.00	0.04	0.00	0.05	0.00	0.08	0.12	0.63
<i>Guidance</i>	86,741	0.07	0.26	0.00	0.00	0.00	0.06	0.00	0.08	0.00	0.05	0.00	0.00	0.02	0.00
<i>PosGuidance</i>	86,741	0.05	0.21	0.00	0.00	0.00	0.04	0.00	0.05	0.00	0.03	0.00	0.00	0.11	0.00
<i>NegGuidance</i>	86,741	0.03	0.18	0.00	0.00	0.00	0.03	0.00	0.04	0.00	0.02	0.00	0.87	0.00	0.00
<i>Edgar8K</i>	86,741	0.28	0.45	0.00	0.00	1.00	0.30	0.00	0.28	0.00	0.28	0.00	0.05	0.21	0.06
<i>Volatility</i>	86,741	0.02	0.03	0.00	0.01	0.02	0.03	0.02	0.02	0.01	0.02	0.01	0.00	0.00	0.00

<i>AbRet_{i,t-60,t-3}</i>	86,741	0.30	18.37	-8.87	-0.53	7.78	3.23	0.44	0.17	-0.43	0.32	-0.83	0.00	0.00	0.29
<i>AbRet_{i,t-2}</i>	86,741	0.02	2.39	-0.88	-0.02	0.86	-0.10	-0.14	0.02	-0.01	0.02	-0.01	0.06	0.08	0.65
<i>AbRet_{i,t-1}</i>	86,741	0.05	2.79	-0.89	-0.01	0.91	-0.10	-0.12	0.04	-0.01	0.08	0.00	0.08	0.02	0.05
<i>Size</i>	86,741	15.87	2.22	14.33	16.02	17.58	15.14	15.24	15.95	16.12	15.77	15.90	0.00	0.00	0.00
<i>BTM</i>	86,741	0.77	1.65	0.19	0.38	0.73	0.42	0.16	0.79	0.38	0.76	0.38	0.00	0.00	0.01
<i>InstOwn</i>	86,741	0.49	0.35	0.02	0.61	0.78	0.53	0.66	0.51	0.64	0.44	0.55	0.00	0.00	0.00
<i>AnalystFollowers</i>	86,741	2.51	0.94	2.08	2.77	3.18	2.30	2.48	2.54	2.83	2.45	2.71	0.00	0.00	0.00
<i>SAFollowers</i>	86,741	6.94	1.96	5.67	7.19	8.33	6.62	6.87	6.92	7.11	7.02	7.39	0.00	0.00	0.00

Table 3 presents descriptive statistics for primary variables used in this study. Each observation represents a unique firm-trading day combination. All variables are defined in Appendix A. Variables marked with “*” are multiplied by 100 for presentation purposes.

Table 5: Test of H1*Panel A: Test of H1*

	(1)	(2)	(3)	(4)	(5)	(6)	(7)
<i>Variable</i>	ALL	ALL	ALL	NO DOW-JONES	NO DOW-JONES OR EARNINGS SURPRISE	NO POST-ARTICLE EVENTS	SINGLE ARTICLE DAYS
<i>Short</i>	-1.045*** (-9.12)	-1.154*** (-9.73)	-1.158*** (-9.74)	-1.552*** (-9.55)	-1.643*** (-9.61)	-1.441*** (-9.47)	-1.173*** (-9.64)
<i>Long</i>	0.431*** (11.92)	0.373*** (11.25)	0.371*** (11.27)	0.505*** (9.77)	0.518*** (9.95)	0.512*** (11.71)	0.362*** (10.76)
<i>NegPct</i>	-16.430*** (-8.52)	-14.452*** (-7.77)	-14.278*** (-7.92)	-15.509*** (-5.75)	-15.781*** (-5.66)	-18.704*** (-7.16)	-12.860*** (-6.95)
<i>PosPct</i>	12.151*** (6.27)	11.197*** (5.49)	10.944*** (5.38)	12.890*** (4.86)	12.867*** (4.88)	11.725*** (4.51)	10.290*** (4.89)
<i>CogProc</i>	-1.615*** (-4.07)	-1.575*** (-3.96)	-1.589*** (-4.00)	-1.327** (-2.24)	-1.160** (-2.01)	-1.616*** (-3.19)	-1.596*** (-4.20)
<i>Numbers</i>	0.542 (1.42)	0.398 (1.01)	0.377 (0.97)	1.121 (1.51)	1.108 (1.45)	-0.010 (-0.02)	0.547 (1.36)
<i>IWordCount</i>	0.037 (1.50)	0.089*** (3.63)	0.091*** (3.73)	0.150*** (4.01)	0.146*** (3.82)	0.166*** (4.98)	0.093*** (3.11)
<i>ComNegPct_{i,t,t+1}</i>	-7.117*** (-5.24)	-5.619*** (-4.26)	-5.524*** (-4.16)	-5.017*** (-2.84)	-4.152** (-2.33)	-5.396*** (-3.04)	-6.275*** (-4.84)
<i>ComPosPct_{i,t,t+1}</i>	2.754*** (2.64)	3.246*** (3.18)	3.190*** (3.10)	4.058*** (2.73)	4.411*** (2.92)	4.545*** (3.52)	3.860*** (3.57)
<i>DJNegPct</i>			-6.800*** (-3.39)			0.237 (0.09)	-5.101** (-2.45)
<i>DJPosPct</i>			17.137*** (8.00)			10.833*** (4.13)	15.512*** (6.88)
<i>IDJ</i>			0.046 (1.39)			0.060 (1.45)	0.045 (1.25)
<i>Upgrades</i>		0.528*** (10.66)	0.522*** (10.55)	0.254*** (2.76)	0.248** (2.38)	0.445*** (6.94)	0.567*** (10.28)
<i>Downgrades</i>		-0.447*** (-10.94)	-0.446*** (-11.08)	-0.235*** (-4.01)	-0.151** (-2.39)	-0.350*** (-5.47)	-0.463*** (-9.99)
<i>ReviseUps</i>		0.013*** (3.27)	0.013*** (3.26)	0.015* (1.93)	0.032*** (3.07)	0.013* (1.67)	0.019*** (3.85)
<i>ReviseDowns</i>		-0.030*** (-7.81)	-0.030*** (-7.71)	-0.013* (-1.85)	-0.014* (-1.76)	-0.025*** (-2.97)	-0.025*** (-5.43)
<i>PosES</i>		0.496*** (6.48)	0.482*** (6.34)	0.061 (0.55)		0.238* (1.78)	0.408*** (4.97)
<i>NegES</i>		-0.903*** (-10.98)	-0.915*** (-11.05)	-0.562*** (-4.64)		-0.560*** (-3.39)	-0.927*** (-9.53)
<i>Guidance</i>		-0.088 (-0.39)	-0.095 (-0.42)	0.155 (0.47)	0.292 (0.63)	-0.033 (-0.12)	-0.259 (-1.07)
<i>PosGuidance</i>		0.454** (2.21)	0.452** (2.20)	-0.008 (-0.03)	-0.159 (-0.36)	0.213 (0.80)	0.565*** (2.64)
<i>NegGuidance</i>		-0.764*** (-3.40)	-0.758*** (-3.37)	-0.452* (-1.70)	-0.798* (-1.90)	-0.506* (-1.91)	-0.565** (-2.56)
<i>Edgar8K</i>		-0.024 (-0.63)	-0.037 (-0.97)	0.018 (0.32)	-0.049 (-0.95)	-0.053 (-1.01)	-0.019 (-0.49)

<i>Volatility</i>		-0.575 (-0.63)	-0.635 (-0.70)	-3.124*** (-2.93)	-2.865** (-2.57)	-0.910 (-0.87)	-0.657 (-0.76)
<i>AbRet_{i,t-60,t-3}</i>		-0.598*** (-5.28)	-0.606*** (-5.35)	-0.313** (-1.98)	-0.302* (-1.77)	-0.688*** (-4.36)	-0.624*** (-5.89)
<i>AbRet_{i,t-2}</i>		-3.895*** (-5.24)	-3.907*** (-5.26)	-2.299** (-2.04)	-2.711** (-2.21)	-4.211*** (-3.78)	-3.640*** (-4.19)
<i>AbRet_{i,t-1}</i>		-3.271*** (-4.62)	-3.323*** (-4.71)	-1.126 (-1.21)	-1.607 (-1.44)	-2.863*** (-3.07)	-2.608*** (-3.28)
<i>Size</i>		-0.077*** (-7.54)	-0.088*** (-7.25)	-0.132*** (-7.88)	-0.136*** (-7.70)	-0.118*** (-8.06)	-0.094*** (-7.67)
<i>BTM</i>		-0.003 (-0.37)	-0.005 (-0.52)	0.002 (0.13)	0.002 (0.12)	-0.008 (-0.80)	-0.014 (-1.34)
<i>InstOwn</i>		0.032 (0.79)	0.029 (0.71)	0.033 (0.64)	-0.004 (-0.08)	-0.018 (-0.32)	-0.026 (-0.63)
<i>AnalystFollowers</i>		-0.072*** (-3.00)	-0.074*** (-3.05)	-0.091*** (-2.70)	-0.084** (-2.42)	-0.082*** (-2.74)	-0.082*** (-3.19)
<i>SAFollowers</i>		0.009 (1.35)	0.008 (1.28)	0.003 (0.29)	-0.002 (-0.22)	0.011 (1.27)	0.010 (1.55)
Observations	86,741	86,741	86,741	37,291	33,641	45,612	74,775
Adjusted R ²	0.011	0.030	0.031	0.027	0.028	0.029	0.029

Panel B: Test of H1 (Article-level, NPA Fixed Effects)

	(1)	(2)	(3)	(4)
<i>Variable</i>	ALL	ALL	NO DOW-JONES	NO DOW-JONES OR EARNINGS SURPRISE
<i>Short</i>	-0.605*** (-5.03)	-0.609*** (-5.06)	-0.841*** (-4.43)	-0.982*** (-4.53)
<i>Long</i>	0.296*** (7.25)	0.295*** (7.29)	0.343*** (5.31)	0.350*** (4.95)
<i>NegPct</i>	-16.983*** (-8.70)	-16.349*** (-8.63)	-15.839*** (-5.40)	-14.781*** (-4.77)
<i>PosPct</i>	10.067*** (4.60)	9.627*** (4.42)	7.284** (2.60)	8.105*** (2.76)
<i>CogProc</i>	-1.602*** (-3.47)	-1.586*** (-3.43)	-1.492** (-2.09)	-1.226 (-1.64)
<i>Numbers</i>	0.966* (1.83)	0.922* (1.75)	1.561 (1.57)	1.546 (1.44)
<i>IWordCount</i>	0.151*** (3.23)	0.153*** (3.31)	0.262*** (3.83)	0.228*** (3.16)
<i>ComPosPct_{i,t,t+1}</i>	2.465** (2.61)	2.318** (2.44)	3.388** (2.30)	3.215** (2.16)
<i>ComNegPct_{i,t,t+1}</i>	-5.963*** (-4.05)	-5.741*** (-3.91)	-3.458* (-1.80)	-2.732 (-1.43)
<i>Constant</i>	1.329 (1.29)	1.336 (1.28)	-1.257 (-1.07)	-0.632 (-0.66)
Observations	104,952	104,952	40,293	36,005
Adjusted R ²	0.062	0.063	0.081	0.085

Table 5 presents results from estimating [1]. The dependent variable is $AbRet_{i,t,t+1}$, multiplied by 100 to facilitate exposition. Panel A reports results using the full-sample of firm-day observations, and Panel B supplements [1] with NPA fixed effects and presents results at the article (rather than firm-day) level. Columns 1 through 3 (1 and 2) include all observations, and Columns 4 and 5 (3 and 4) exclude observations with concurrently issued Dow-Jones news content and Dow Jones news content and earnings surprises in Panel A (Panel B). Columns 6 and 7 in Panel A also exclude articles with post-article information events and articles on days on which more than one article is published. All variables are defined in Appendix A. All estimations include year-month and industry fixed effects. *** (**, *) denotes two-tailed significance at the $p < 0.01$ ($p < 0.05$, $p < 0.10$) level (two-tailed) assessed using t -statistics (in parentheses) derived from White (1980) standard errors clustered by year-month to correct for cross-sectional correlation in error terms.

Table 6: Test of H2

<i>Variable</i>	(1)	(2)	(3)	(4)	<i>Test of Difference</i>			
	<i>Position=0</i>	<i> Position =1</i>	<i>Position=-1</i>	<i>Position=1</i>	1-2	1-3	1-4	3-4
<i>NegPct</i>	-11.696*** (-5.88)	-31.925*** (-9.17)	-19.280* (-1.77)	-19.424*** (-4.91)	0.00	0.24	0.04	0.99
<i>PosPct</i>	8.862*** (4.52)	23.300*** (5.52)	37.616* (1.97)	12.341*** (3.10)	0.00	0.06	0.20	0.17
<i>CogProc</i>	-1.346*** (-3.18)	-2.642*** (-2.84)	-1.580 (-0.49)	-2.429*** (-2.84)	0.16	0.91	0.19	0.78
<i>Numbers</i>	0.854* (1.84)	-0.443 (-0.57)	-4.301 (-1.26)	0.430 (0.56)	0.17	0.13	0.65	0.16
<i>lWordCount</i>	0.066** (2.12)	0.081** (2.06)	-0.385** (-2.50)	0.185*** (4.74)	0.76	0.00	0.01	0.00
<i>ComNegPct_{i,t,t+1}</i>	-4.828*** (-3.72)	-10.638*** (-3.87)	4.903 (0.45)	-8.709*** (-3.14)	0.04	0.36	0.17	0.22
<i>ComPosPct_{i,t,t+1}</i>	1.937 (1.47)	8.275*** (3.47)	-8.084 (-0.64)	6.358*** (2.82)	0.03	0.43	0.11	0.22
<i>DJNegPct</i>	-9.387*** (-4.16)	-1.320 (-0.41)	-22.341* (-1.84)	-0.450 (-0.13)	0.03	0.25	0.03	0.08
<i>DJPosPct</i>	16.411*** (6.86)	18.176*** (4.52)	17.408 (0.85)	18.205*** (4.35)	0.73	0.97	0.75	0.97
<i>IDJ</i>	0.048 (1.26)	0.020 (0.31)	0.080 (0.29)	0.042 (0.62)	0.75	0.89	0.98	0.89
<i>Upgrades</i>	0.535*** (10.39)	0.491*** (6.72)	0.349 (1.12)	0.490*** (6.56)	0.55	0.54	0.54	0.65
<i>Downgrades</i>	-0.434*** (-8.94)	-0.467*** (-7.57)	-0.710** (-2.56)	-0.436*** (-6.90)	0.63	0.31	0.96	0.31
<i>ReviseUps</i>	0.012** (2.14)	0.018*** (2.78)	-0.008 (-0.18)	0.018*** (3.09)	0.51	0.63	0.45	0.52
<i>ReviseDowns</i>	-0.027*** (-5.42)	-0.036*** (-5.36)	-0.087** (-2.54)	-0.035*** (-5.44)	0.23	0.07	0.30	0.12
<i>PosES</i>	0.511*** (6.02)	0.405*** (2.84)	1.710** (2.39)	0.314** (2.09)	0.52	0.08	0.26	0.05
<i>NegES</i>	-1.090*** (-9.59)	-0.542*** (-3.20)	1.465** (2.24)	-0.737*** (-4.01)	0.02	0.00	0.13	0.00
<i>Guidance</i>	-0.287 (-1.22)	0.455 (0.78)	3.098* (1.85)	-0.073 (-0.15)	0.25	0.04	0.72	0.04
<i>PosGuidance</i>	0.572*** (2.78)	0.090 (0.16)	-2.512* (-1.71)	0.638 (1.31)	0.45	0.03	0.89	0.01
<i>NegGuidance</i>	-0.660*** (-2.75)	-1.105** (-2.12)	-4.733*** (-2.94)	-0.437 (-1.03)	0.44	0.01	0.65	0.00
<i>Edgar8K</i>	0.017 (0.40)	-0.138** (-2.12)	-0.644* (-1.78)	-0.091 (-1.17)	0.04	0.07	0.21	0.15
<i>Volatility</i>	-1.813 (-1.60)	0.179 (0.13)	-5.325* (-1.93)	3.229** (2.29)	0.25	0.22	0.00	0.00
<i>AbRet_{i,t-60,t-3}</i>	-0.483*** (-3.52)	-0.987*** (-5.93)	-1.700*** (-3.67)	-0.655*** (-3.37)	0.01	0.01	0.37	0.04
<i>AbRet_{i,t-2}</i>	-4.205*** (-3.83)	-3.587*** (-2.68)	-12.895*** (-3.43)	-1.872 (-1.31)	0.76	0.03	0.24	0.00
<i>AbRet_{i,t-1}</i>	-3.549*** (-3.92)	-3.242** (-2.31)	-12.769*** (-3.72)	-2.273 (-1.53)	0.88	0.01	0.50	0.00

<i>Size</i>	-0.051*** (-3.90)	-0.141*** (-6.85)	0.358*** (3.56)	-0.166*** (-7.47)	0.00	0.00	0.00	0.00
<i>BTM</i>	0.013 (1.15)	-0.020 (-1.11)	0.066 (0.42)	-0.054*** (-3.41)	0.12	0.72	0.00	0.42
<i>InstOwn</i>	-0.013 (-0.27)	0.051 (0.72)	0.467 (1.11)	0.076 (0.95)	0.49	0.24	0.36	0.36
<i>AnalystFollowers</i>	-0.012 (-0.43)	-0.171*** (-3.89)	0.430** (2.22)	-0.287*** (-6.99)	0.00	0.02	0.00	0.00
<i>SAFollowers</i>	0.006 (0.71)	0.034*** (2.62)	-0.009 (-0.17)	0.039*** (2.95)	0.06	0.81	0.03	0.39
Observations	58,057	28,684	2,431	26,253				
Adjusted R ²	0.025	0.036	0.108	0.044				

Table 6 presents results from estimating [1] separately by position and tests of differences in coefficients across the four columns. The dependent variable is $AbRet_{i,t,t+1}$, multiplied by 100 to facilitate exposition. All variables are defined in Appendix A. All estimations include year-month and industry fixed effects. *** (**, *) denotes significance at the $p < 0.01$ ($p < 0.05$, $p < 0.10$) level (one-tailed where a prediction is made and two-tailed otherwise) assessed using t -statistics (in parentheses) derived from White (1980) standard errors clustered by year-month to correct for cross-sectional correlation in error terms.

Table 7: Early Morning SA Articles

	(1)	(2)	(3)	(4)	(5)	(6)	(7)
<i>Variable</i>	ALL	ALL	ALL	NO DOW- JONES	NO DOW- JONES OR EARNINGS SURPRISE	NO POST ARTICLE EVENTS	SINGLE ARTICLE DAYS
<i>Short</i>	-1.312*** (-6.86)	-1.398*** (-6.97)	-1.403*** (-6.99)	-1.896*** (-6.46)	-2.051*** (-6.16)	-1.385*** (-5.46)	-1.462*** (-6.89)
<i>Long</i>	0.371*** (6.12)	0.323*** (5.33)	0.321*** (5.32)	0.483*** (4.58)	0.524*** (4.74)	0.550*** (5.74)	0.319*** (5.17)
<i>NegPct</i>	-7.247* (-1.87)	-6.655* (-1.66)	-6.692* (-1.67)	-15.825** (-2.51)	-18.039*** (-2.78)	-15.703*** (-2.77)	-6.413 (-1.57)
<i>PosPct</i>	8.585** (2.31)	9.181** (2.46)	9.147** (2.44)	15.420** (2.49)	15.221** (2.43)	13.485** (2.43)	9.710** (2.60)
<i>CogProc</i>	-1.721** (-2.29)	-1.472* (-1.93)	-1.471* (-1.93)	-0.005 (-0.00)	-0.210 (-0.18)	-0.919 (-0.91)	-1.475* (-1.92)
<i>Numbers</i>	1.275 (1.58)	1.259 (1.50)	1.250 (1.50)	1.314 (0.95)	0.945 (0.66)	-0.331 (-0.28)	1.398* (1.69)
<i>lWordCount</i>	0.097* (1.86)	0.079 (1.49)	0.079 (1.51)	0.054 (0.75)	0.056 (0.74)	0.145** (2.12)	0.083 (1.53)
<i>ComNegPct_{i,t,t+1}</i>	-5.105** (-2.09)	-4.071 (-1.65)	-3.990 (-1.61)	-1.289 (-0.37)	-1.113 (-0.31)	-6.123** (-2.07)	-4.319* (-1.76)
<i>ComPosPct_{i,t,t+1}</i>	3.058 (1.51)	3.363 (1.66)	3.376* (1.66)	3.679 (1.30)	4.080 (1.36)	4.909* (1.68)	3.294 (1.60)
<i>Upgrades</i>		0.286*** (2.90)	0.284*** (2.87)	0.174 (1.00)	0.297 (1.53)	-0.014 (-0.09)	0.265*** (2.69)
<i>Downgrades</i>		-0.189** (-2.24)	-0.187** (-2.22)	-0.213 (-1.53)	-0.124 (-0.82)	-0.129 (-0.98)	-0.207** (-2.32)
<i>Volatility</i>		-2.012 (-1.09)	-2.090 (-1.14)	1.088 (0.48)	0.838 (0.37)	-0.249 (-0.12)	-1.700 (-0.92)
<i>AbRet_{i,t-60,t-3}</i>		-0.658*** (-3.00)	-0.660*** (-2.99)	-0.692** (-2.40)	-0.656** (-2.10)	-0.756** (-2.59)	-0.689*** (-3.09)
<i>AbRet_{i,t-2}</i>		-2.317 (-1.55)	-2.325 (-1.56)	-0.256 (-0.11)	-2.416 (-1.01)	-3.225 (-1.48)	-2.289 (-1.48)
<i>AbRet_{i,t-1}</i>		-1.880 (-1.33)	-1.911 (-1.35)	-1.408 (-0.70)	-2.485 (-1.04)	-1.509 (-0.77)	-1.735 (-1.20)
<i>PosES</i>		0.215 (1.46)	0.205 (1.39)	0.095 (0.43)		0.268 (1.19)	0.238 (1.52)
<i>NegES</i>		-0.704*** (-4.19)	-0.711*** (-4.26)	-0.542* (-1.96)		-0.639** (-2.24)	-0.713*** (-4.03)
<i>DJNegPct</i>			-1.980 (-0.69)			-2.890 (-0.62)	-3.075 (-0.99)
<i>DJPosPct</i>			10.002*** (2.99)			4.488 (0.98)	11.686*** (3.49)
<i>IDJ</i>			-0.008 (-0.11)			-0.028 (-0.28)	-0.034 (-0.47)
<i>ReviseUps</i>		0.005 (0.54)	0.005 (0.54)	-0.001 (-0.06)	0.003 (0.14)	0.013 (0.92)	0.010 (1.12)
<i>ReviseDowns</i>		0.004 (0.50)	0.004 (0.51)	0.023 (1.56)	0.029* (1.97)	0.023 (1.59)	0.003 (0.36)
<i>Guidance</i>		-0.038	-0.048	-0.129	-0.657	-0.583	-0.071

		(-0.09)	(-0.12)	(-0.26)	(-1.07)	(-1.03)	(-0.17)
<i>PosGuidance</i>		0.267	0.269	-0.109	0.640	0.398	0.162
		(0.72)	(0.73)	(-0.24)	(1.07)	(0.78)	(0.42)
<i>NegGuidance</i>		-0.494	-0.482	0.009	0.391	0.252	-0.449
		(-1.31)	(-1.28)	(0.02)	(0.67)	(0.54)	(-1.14)
<i>Edgar8K</i>		-0.067	-0.076	-0.059	-0.093	-0.166	-0.058
		(-0.85)	(-0.97)	(-0.50)	(-0.78)	(-1.54)	(-0.73)
<i>Size</i>		-0.054***	-0.061***	-0.058*	-0.063*	-0.053*	-0.054**
		(-2.70)	(-2.81)	(-1.69)	(-1.75)	(-1.82)	(-2.46)
<i>BTM</i>		-0.022	-0.024	-0.016	-0.012	-0.004	-0.025
		(-1.22)	(-1.29)	(-0.56)	(-0.40)	(-0.18)	(-1.36)
<i>InstOwn</i>		0.170**	0.171**	0.095	0.102	0.055	0.187***
		(2.44)	(2.46)	(0.92)	(0.92)	(0.56)	(2.84)
<i>AnalystFollowers</i>		-0.156***	-0.156***	-0.273***	-0.267***	-0.159***	-0.165***
		(-3.16)	(-3.16)	(-4.15)	(-4.00)	(-2.73)	(-3.40)
<i>SAFollowers</i>		0.007	0.007	0.011	0.004	0.008	0.007
		(0.68)	(0.68)	(0.59)	(0.18)	(0.50)	(0.63)
Observations	18,782	18,782	18,782	9,042	8,218	10,259	18,141
Adjusted R ²	0.015	0.023	0.023	0.033	0.035	0.028	0.024

Table 7 presents results from estimating [1] using a sample of articles published between 9:30 am and 1:00 pm on trading days. The dependent variable is the two-day abnormal return, where the day 0 return is defined as (closing price – opening price) / opening price. The two-day abnormal return is multiplied by 100 to facilitate exposition. Columns 1 through 3 impose no additional sample screens, and Column 4 (5) [6] {7} excludes observations with concurrently issued Dow-Jones news content (Dow Jones news content or earnings surprises) [post-article information events] {articles on days on which more than one article is written}. All variables are defined in Appendix A. All estimations include year-month and industry fixed effects. *** (**, *) denotes two-tailed significance at the $p < 0.01$ ($p < 0.05$, $p < 0.10$) level (two-tailed) assessed using t -statistics (in parentheses) derived from White (1980) standard errors clustered by year-month to correct for cross-sectional correlation in error terms.

Table 8: NPA Positions and Article Length, Rigor, and Numerical Content

	(1)	(2)	(3)	(4)
<i>Variable</i>	ALL	NO DOW-JONES	NO DOW-JONES OR EARNINGS SURPRISE	NO POST EVENTS
<i>Short</i>	-0.988*** (-7.94)	-1.267*** (-7.71)	-1.354*** (-7.63)	-1.210*** (-7.49)
<i>Long</i>	0.368*** (11.15)	0.493*** (9.51)	0.503*** (9.62)	0.485*** (10.63)
<i>Short x CogProc</i>	-1.496 (-0.47)	0.076 (0.02)	0.296 (0.08)	-1.077 (-0.26)
<i>Short x Numbers</i>	-7.529** (-2.18)	-8.325* (-1.72)	-9.024* (-1.86)	-3.907 (-0.89)
<i>Short x lWordCount</i>	-0.899*** (-5.85)	-0.965*** (-4.36)	-0.977*** (-4.41)	-0.833*** (-3.69)
<i>Long x CogProc</i>	-0.168 (-0.18)	2.461 (1.64)	2.340 (1.53)	-1.049 (-0.69)
<i>Long x Numbers</i>	-0.449 (-0.51)	-1.125 (-0.89)	-0.731 (-0.51)	1.512 (1.21)
<i>Long x lWordCount</i>	0.165*** (3.24)	0.250*** (3.02)	0.236*** (2.84)	0.264*** (3.66)
<i>NegPct</i>	-14.400*** (-7.98)	-15.634*** (-5.83)	-15.829*** (-5.70)	-19.688*** (-7.07)
<i>PosPct</i>	11.142*** (5.51)	13.245*** (5.04)	13.223*** (5.03)	12.963*** (4.70)
<i>CogProc</i>	-1.448*** (-3.40)	-1.925*** (-2.70)	-1.740** (-2.56)	-0.992* (-1.75)
<i>Numbers</i>	0.628 (1.40)	1.578** (2.01)	1.492* (1.81)	-0.130 (-0.21)
<i>lWordCount</i>	0.072** (2.46)	0.108** (2.25)	0.110** (2.29)	0.132*** (3.21)
<i>ComNegPct_{i,t,t+1}</i>	-5.334*** (-4.06)	-4.863*** (-2.76)	-4.027** (-2.26)	-4.031** (-2.27)
<i>ComPosPct_{i,t,t+1}</i>	3.108*** (3.02)	4.071*** (2.73)	4.404*** (2.89)	4.610*** (3.56)
<i>DJNegPct</i>	-6.752*** (-3.36)			1.322 (0.45)
<i>DJPosPct</i>	17.145*** (8.00)			11.803*** (4.23)
<i>IDJ</i>	0.046 (1.40)			0.063 (1.51)
<i>Upgrades</i>	0.524*** (10.56)	0.254*** (2.76)	0.248** (2.37)	0.449*** (6.18)
<i>Downgrades</i>	-0.445*** (-11.05)	-0.231*** (-3.94)	-0.146** (-2.32)	-0.313*** (-4.81)
<i>ReviseUps</i>	0.013*** (3.23)	0.015* (1.91)	0.032*** (3.04)	0.008 (0.97)
<i>ReviseDowns</i>	-0.030*** (-7.74)	-0.013* (-1.87)	-0.014* (-1.78)	-0.026*** (-3.02)
<i>PosES</i>	0.483*** (6.36)	0.065 (0.59)		0.210 (1.46)
<i>NegES</i>	-0.919***	-0.563***		-0.605***

	(-11.06)	(-4.62)		(-3.30)
<i>Guidance</i>	-0.108 (-0.48)	0.141 (0.43)	0.279 (0.60)	0.031 (0.11)
<i>PosGuidance</i>	0.464** (2.28)	-0.000 (-0.00)	-0.156 (-0.35)	0.232 (0.88)
<i>NegGuidance</i>	-0.750*** (-3.35)	-0.447* (-1.70)	-0.797* (-1.91)	-0.483* (-1.93)
<i>Edgar8K</i>	-0.035 (-0.92)	0.020 (0.36)	-0.047 (-0.90)	-0.065 (-1.21)
<i>Volatility</i>	-0.548 (-0.61)	-3.069*** (-2.89)	-2.806** (-2.52)	-1.351 (-1.21)
<i>AbRet_{i,t-60,t-3}</i>	-0.608*** (-5.37)	-0.314** (-1.99)	-0.303* (-1.78)	-0.540*** (-3.23)
<i>AbRet_{i,t-2}</i>	-3.939*** (-5.31)	-2.380** (-2.11)	-2.817** (-2.28)	-4.873*** (-4.25)
<i>AbRet_{i,t-1}</i>	-3.346*** (-4.75)	-1.175 (-1.26)	-1.650 (-1.48)	-3.167*** (-3.49)
<i>Size</i>	-0.089*** (-7.31)	-0.132*** (-7.85)	-0.137*** (-7.68)	-0.119*** (-8.04)
<i>BTM</i>	-0.006 (-0.59)	0.002 (0.17)	0.002 (0.17)	-0.016 (-1.46)
<i>InstOwn</i>	0.034 (0.82)	0.037 (0.70)	-0.002 (-0.03)	-0.011 (-0.19)
<i>AnalystFollowers</i>	-0.078*** (-3.26)	-0.094*** (-2.82)	-0.087** (-2.51)	-0.112*** (-3.78)
<i>SAFollowers</i>	0.008 (1.25)	0.004 (0.38)	-0.001 (-0.13)	0.009 (0.97)
Observations	86,741	37,291	33,641	41,075
Adjusted R ²	0.032	0.028	0.029	0.031

Table 8 presents results from estimating [1] after interacting article characteristics with NPA positions. Column 1 includes all observations, and Column 2 (3) [4] excludes firm-days with contemporaneously issued Dow-Jones content (contemporaneously issued Dow-Jones content or earnings-surprise announcements) [post-article information events]. All variables are defined in Appendix A. All estimations include year-month and industry fixed effects. *** (**, *) denotes significance at the $p < 0.01$ ($p < 0.05$, $p < 0.10$) level (one-tailed where a prediction is made and two-tailed otherwise two-tailed) assessed using t -statistics (in parentheses) derived from White (1980) standard errors clustered by year-month to correct for cross-sectional correlation in error terms.

Table 9: First-time vs. Repeated Disclosures of Position

	(1)	(2)	(3)	(4)
	ALL	NO DJ	NO DJ OR EARN OR EARNINGS SURPRISE	NO POST EVENT
<i>Variable</i>				
<i>FirstDisc x Short</i>	-1.314*** (-5.83)	-1.668*** (-4.50)	-1.701*** (-4.51)	-1.445*** (-4.96)
<i>FirstDisc x Long</i>	0.205*** (3.43)	0.312*** (3.29)	0.296*** (3.11)	0.292*** (3.21)
<i>FirstDisc x NegPct</i>	-0.604 (-0.18)	-4.345 (-0.83)	-5.827 (-1.14)	6.092 (1.18)
<i>FirstDisc x PosPct</i>	-6.992* (-1.81)	-6.647 (-1.17)	-9.666* (-1.66)	-7.758 (-1.63)
<i>FirstDisc x CogProc</i>	2.925*** (3.54)	1.934 (1.63)	1.188 (0.91)	1.654 (1.55)
<i>FirstDisc x Numbers</i>	-1.625* (-1.68)	-1.212 (-0.96)	-1.528 (-1.15)	-1.682 (-1.39)
<i>FirstDisc x lWordCount</i>	-0.086* (-1.73)	-0.218*** (-2.94)	-0.204*** (-2.64)	-0.141* (-1.79)
<i>FirstDisc x ComPosPct_{i,t,t+1}</i>	0.833 (0.39)	1.989 (0.58)	5.135 (1.49)	1.084 (0.35)
<i>FirstDisc x ComNegPct_{i,t,t+1}</i>	-0.466 (-0.17)	-4.185 (-1.05)	-4.286 (-1.02)	0.668 (0.16)
Observations	86,741	37,291	33,641	41,075
Adjusted R-squared	0.032	0.029	0.030	0.031

Table 9 presents results from estimating [1] after including interactions between SA article-related variables and *FirstDisc*, an indicator equaling 1 the first time the NPA discloses a position. All other control variables from [1] are included, but coefficient estimates are suppressed to facilitate exposition. Column 1 includes all observations, and Column 2 (3) [4] excludes firm-days with contemporaneously issued Dow-Jones content (contemporaneously issued Dow-Jones content or earnings-surprise announcements) [post-article information events]. All variables are defined in Appendix A. All estimations include year-month and industry fixed effects. *** (**, *) denotes significance at the $p < 0.01$ ($p < 0.05$, $p < 0.10$) level (one-tailed where a prediction is made and two-tailed otherwise) assessed using t -statistics (in parentheses) derived from White (1980) standard errors clustered by year-month to correct for cross-sectional correlation in error terms.

Table 10: NPA Position and Post-Publication Drift

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
<i>Variable</i>	ALL 60 Days	ALL 60 Days	ALL 60 Days	NO DOW- JONES 60 Days	NO DOW- JONES OR EARNINGS SURPRISE 60 Days	ALL Day 3-5	ALL Day 3-10	ALL Day 3-20
<i>Short</i>	-2.070*** (-3.51)	-1.510*** (-2.69)	-1.526*** (-2.72)	-1.942** (-2.55)	-1.883** (-2.39)	-0.001 (-1.37)	-0.001 (-0.62)	-0.005 (-1.51)
<i>Long</i>	-0.282 (-1.38)	0.005 (0.03)	-0.006 (-0.03)	-0.192 (-0.67)	-0.257 (-0.84)	0.000 (0.30)	0.001 (1.03)	0.001 (0.88)
<i>NegPct</i>	-4.709 (-0.52)	-6.889 (-0.77)	-9.599 (-1.09)	-20.541 (-1.54)	-17.318 (-1.30)	0.010 (0.60)	-0.000 (-0.02)	-0.039 (-0.95)
<i>PosPct</i>	4.049 (0.43)	-3.462 (-0.38)	-2.233 (-0.25)	6.293 (0.44)	8.244 (0.55)	0.014 (0.78)	0.006 (0.20)	-0.017 (-0.39)
<i>CogProc</i>	-3.638 (-1.54)	-2.218 (-1.01)	-2.387 (-1.09)	-6.216* (-1.87)	-6.958** (-2.13)	-0.004 (-0.83)	-0.010 (-1.29)	-0.016 (-1.23)
<i>Numbers</i>	5.239** (2.19)	3.648 (1.58)	3.951* (1.71)	7.356** (2.10)	7.239** (2.01)	0.006 (1.64)	0.010 (1.28)	0.028** (2.39)
<i>lWordCount</i>	0.345** (2.40)	0.351** (2.49)	0.353** (2.51)	0.428 (1.63)	0.449 (1.58)	0.000 (1.17)	0.001 (1.66)	0.000 (0.54)
<i>ComNegPcti_[t,t+1]</i>	4.226 (0.57)	2.984 (0.43)	2.083 (0.30)	-4.395 (-0.42)	-4.871 (-0.44)	0.000 (0.01)	-0.038* (-1.67)	-0.010 (-0.29)
<i>ComPosPcti_[t,t+1]</i>	-6.293 (-1.38)	-7.515* (-1.67)	-7.159 (-1.59)	-2.819 (-0.34)	-0.883 (-0.10)	0.010 (0.88)	0.015 (0.86)	0.036 (1.25)
<i>ComNegPcti_[t+3,t+60]</i>	-33.368*** (-5.09)	-30.581*** (-4.65)	-30.326*** (-4.62)	-47.694*** (-4.42)	-45.243*** (-4.03)	-0.059*** (-4.45)	-0.135*** (-6.17)	-0.193*** (-5.75)
<i>ComPosPcti_[t+3,t+60]</i>	6.163 (1.38)	8.021* (1.79)	8.208* (1.83)	19.825** (2.49)	24.444*** (2.96)	0.003 (0.25)	0.047*** (2.76)	0.068*** (2.72)
<i>DJNegPct</i>			29.500*** (3.61)			-0.009 (-0.52)	0.051* (1.79)	0.088** (2.21)
<i>DJPosPct</i>			-20.735** (-2.03)			-0.018 (-0.86)	-0.052 (-1.62)	-0.083 (-1.54)
<i>IDJ</i>			0.243 (1.26)			0.000 (0.56)	0.000 (0.65)	0.002 (1.34)
<i>Upgrades</i>		0.104 (0.44)	0.084 (0.36)	0.457 (1.04)	0.351 (0.62)	0.000 (0.61)	0.000 (0.30)	0.000 (0.21)

<i>Downgrades</i>		-0.422*	-0.451**	-0.613*	-0.877**	0.000	-0.001	-0.001
		(-1.97)	(-2.09)	(-1.72)	(-2.30)	(0.47)	(-1.46)	(-0.71)
<i>ReviseUps</i>		0.046*	0.045*	0.086**	0.097*	0.000	0.000	0.000
		(1.92)	(1.89)	(2.11)	(1.79)	(0.89)	(0.68)	(1.64)
<i>ReviseDowns</i>		-0.061***	-0.065***	-0.095**	-0.079	-0.000	-0.000	-0.000
		(-3.03)	(-3.24)	(-2.16)	(-1.47)	(-1.32)	(-0.64)	(-1.07)
<i>PosES</i>		0.468	0.427	0.390		0.000	0.002*	0.002
		(1.44)	(1.32)	(0.69)		(0.77)	(1.80)	(1.22)
<i>NegES</i>		0.045	0.001	-0.662		-0.000	0.000	-0.002
		(0.13)	(0.00)	(-1.24)		(-0.03)	(0.24)	(-0.83)
<i>Guidance</i>		-0.789	-0.774	1.438	-0.377	0.002	0.002	0.002
		(-0.96)	(-0.94)	(1.22)	(-0.20)	(1.20)	(0.72)	(0.52)
<i>PosGuidance</i>		0.771	0.757	-1.081	1.244	-0.001	-0.002	-0.001
		(0.96)	(0.94)	(-0.97)	(0.67)	(-0.89)	(-0.78)	(-0.17)
<i>NegGuidance</i>		0.045	0.016	-1.771*	-1.002	-0.001	-0.001	-0.005
		(0.06)	(0.02)	(-1.69)	(-0.61)	(-0.85)	(-0.73)	(-1.48)
<i>Edgar8K</i>		-0.068	-0.113	0.256	0.223	-0.000	0.000	0.000
		(-0.32)	(-0.53)	(0.87)	(0.81)	(-1.17)	(0.26)	(0.05)
<i>Volatility</i>		-13.394*	-14.165*	-21.883**	-20.111**	-0.038***	-0.061***	-0.092***
		(-1.86)	(-1.97)	(-2.47)	(-2.28)	(-3.79)	(-3.34)	(-2.69)
<i>AbReti_{i,t-60,t-3}</i>		-1.632	-1.618	-2.398*	-2.475**	-0.003**	-0.006**	-0.009
		(-1.38)	(-1.37)	(-1.95)	(-1.99)	(-1.98)	(-2.23)	(-1.43)
<i>AbReti_{i,t-2}</i>		-4.899	-4.917	-5.308	-5.864	-0.014*	-0.027**	-0.034*
		(-1.22)	(-1.22)	(-1.08)	(-0.98)	(-1.75)	(-2.11)	(-1.76)
<i>AbReti_{i,t-1}</i>		-7.127**	-7.049**	-10.192**	-11.344**	-0.010	-0.017	-0.051***
		(-2.32)	(-2.30)	(-2.58)	(-2.37)	(-1.54)	(-1.50)	(-3.38)
<i>AbReti_{i,t,t+1}</i>	-0.016	-0.022	-0.021	-0.036	-0.061*	-0.000	-0.000	-0.000*
	(-0.64)	(-0.90)	(-0.86)	(-1.06)	(-1.73)	(-0.52)	(-0.96)	(-1.81)
<i>Size</i>		0.144	0.086	0.093	0.110	-0.000	-0.000	0.000
		(1.66)	(0.93)	(0.98)	(1.14)	(-0.74)	(-0.44)	(0.12)
<i>BTM</i>		0.622***	0.599***	0.543***	0.501***	0.000*	0.001***	0.002***
		(8.53)	(8.17)	(4.93)	(4.61)	(1.79)	(3.98)	(5.61)
<i>InstOwn</i>		1.961***	1.976***	2.298***	2.373***	0.001**	0.003***	0.006***
		(4.55)	(4.59)	(5.54)	(5.60)	(2.54)	(3.31)	(2.91)
<i>AnalystFollowers</i>		0.229	0.239	-0.110	-0.127	0.000	0.001**	0.001
		(0.90)	(0.94)	(-0.43)	(-0.49)	(1.45)	(2.11)	(1.24)
<i>SAFollowers</i>		0.010	0.007	0.042	0.015	-0.000	-0.000	-0.000
		(0.26)	(0.19)	(0.79)	(0.26)	(-0.45)	(-1.07)	(-1.10)

Observations	86,641	86,641	86,641	37,221	33,575	86,731	86,714	86,697
Adjusted R ²	0.022	0.029	0.029	0.031	0.031	0.004	0.008	0.013

Table 10 presents results from estimating [1] using $AbRet_{i,[t+3,t+60]}$ (multiplied by 100) as the dependent variable. Columns 1 to 3 include all observations, and Column 4 (5) excludes observations with concurrently issued Dow-Jones news content (Dow Jones news content or earnings surprises). Columns 5-8 repeat the analyses in Columns 1-3 for drift during the 3 to 5 days, 3 to 10 days and 3 to 20 days following the article date. All variables are defined in Appendix A. All estimations include year-month and industry fixed effects. *** (**, *) denotes significance at the $p < 0.01$ ($p < 0.05$, $p < 0.10$) level (one-tailed where a prediction is made and two-tailed otherwise) assessed using t -statistics (in parentheses) derived from White (1980) standard errors clustered by year-month to correct for cross-sectional correlation in error terms.